

RUNS OF PRECIPITATION SERIES

by

Jose Llamas and M. M. Siddiqui

May 1969



HYDROLOGY PAPERS
COLORADO STATE UNIVERSITY
Fort Collins, Colorado

RUNS OF PRECIPITATION SERIES

by

Jose Llamas

and

M. M. Siddiqui

HYDROLOGY PAPERS

COLORADO STATE UNIVERSITY

FORT COLLINS, COLORADO 80521

May 1969

No. 33

TABLE OF CONTENTS

	<u>Page</u>
Abstract	1
I Introduction	1
1.1 Subject of this study	1
1.2 Background of the problem	1
II A Sequence of Independent Variables	3
2.1 Introduction.	3
2.2 Formulation of the problem	3
2.3 The independence of $\{(N_{1j}, S_{1j})\}$ and $\{(N_{2j}, S_{2j})\}$	3
2.4 The distribution of S_1 in some special cases	4
2.5 Moment generating functions	4
III Normal and Gamma Distributed Variables	6
3.1 Normal sequence.	6
3.2 Approximated probability density functions of I_1 and I_2	6
3.3 Gamma distributed sequence	7
3.4 Example: Fort Collins, Station No. 5.3005	9
IV Two Mutually Independent Processes	12
4.1 Introduction.	12
4.2 Formulation of the problem	12
4.3 Comment on the dependent case	13
V Monthly Precipitation Series	14
5.1 Introduction	14
5.2 Formulation of the problem	14
5.3 Properties of runs.	15
5.4 Example	15
5.5 Explanation of appendices	15
References.	16
Appendix I	18
Appendix II	20

LIST OF FIGURES AND TABLES

Table

1	Improvement of probability density function of S_i ($i = 1, 2, \dots$)	7
2	Improvement of Probability density function of I_i ($i = 1, 2, \dots$)	7
3	Moments of X_{i1}^* ($i = 1, 2, \dots$) for several values of r	9
4	Moments of S_i ($i = 1, 2, \dots$) for several values of r	9
5	Moments of I_i ($i = 1, 2, \dots$) for several values of r	9
6	Expected values and variances of X_i and N_i ($i = 1, 2, \dots$) for different hypotheses of $\{X_n\}$	11
7	Expected values and variances of S_i and I_i ($i = 1, 2, \dots$) for different hypotheses of $\{X_n\}$	11

ABSTRACT

Three quantitative measures are introduced for the concepts of "surplus" and "deficit" in hydrologic series. These are: run-length, run-sum and run-intensity. Positive and negative runs of a series are defined in terms of a fixed value, say c , of the variable under consideration, namely precipitation. The distribution function, moments, and other statistical properties of the three variables, run-length, run-sum, and run-intensity, are obtained analytically under the following alternative assumptions on the sequence of annual precipitations:

1. It is independent and normally distributed.
2. It is independent and gamma distributed.

For monthly precipitations, z_t , the series was first standardized by the transformation

$$x_t = \frac{z_t - \mu_\tau}{\sigma_\tau} ,$$

where t is of the form $t = 12(n-1) + \tau$, $\tau = 1, \dots, 12$, $n = 1, 2, \dots$, and where μ_τ and σ_τ are mean and standard deviation of the series corresponding to the month τ . Calling " $x_t \leq c$ " as state "0" and " $x_t > c$ " as state "1," the series is then analyzed as a two-state Markov chain with stationary transition probabilities.

Annual precipitation from 27 stations in Colorado, and monthly precipitation from 219 stations in the Western United States are analyzed.

RUNS OF PRECIPITATION SERIES

by

Jose Llamas* and M. M. Siddiqui**

Chapter I

INTRODUCTION

1.1 Subject of this study. The major objective of this study is to carry out the mathematical analysis of some parameters by which the concept of runs of a precipitation series may be defined with reference to the series itself. One of the main problems in water resource projects is to predict accurately the amount of water available during a period of operation and to determine whether or not it will be sufficient. The total amount of water necessary in a given period of time, whether for one particular project or for a number of projects in one region, can be considered as the water demand of that region. Of course, the demand changes from region to region or from country to country. For instance, in arid areas the water demand must be necessarily less than in the humid regions because of different water availability. The same situation is encountered in the agricultural countries as compared to the industrial ones. If in one period of time the supply of water is smaller or greater than the demand, then this period can be considered as dry or wet, respectively.

The concepts of dry or wet periods ought to be taken only in a relative sense so that they depend on a certain level, c . This value, c , can be a constant or a variable according to the characteristics of the water demand. In the case of agricultural projects on a constant surface of land and for the same kind of annual crops, the consumptive use of water is usually constant every year. In the case of urban development, the future requirement of water is related to the growth of the population and to the expected industrial expansion.

1.2 Background of the problem. The problem of runs of a precipitation series has been initiated by Downer, Siddiqui and Yevjevich [1], and Yevjevich [2]. In these two papers, the authors define a dry or a wet run or a negative or positive run as the period in which the total amount of precipitation is less or greater than a certain constant, c . This constant may correspond to the concept of water demands previously defined. Three main factors may be used in order to characterize a particular negative or positive run: run-length, run-sum, and run-intensity. The run-length of a wet (positive) or a dry (negative) run is the number of terms in a complete positive or negative run, respectively. This is also the duration of a positive or a negative run. This quantity is particularly

important in water resource problems because the knowledge of the expected duration of drought or rainfall provides the engineer with the necessary design information.

The run-sum (or the magnitude of a run) is defined as the sum of deviations from a level (water demand) of precipitation over the run-length. These deviations are negative or positive when the run is dry or wet, respectively. In some water resource problems the run-sum is the most important factor. The total capacity of water that must be stored and then supplied depends on the expected run-sum of the future dry negative run. The run-sum of positive or negative runs is directly related to the sizing of reservoir capacities, design and operation of hydroelectric structures, projects of water pollution, sizing of pumps, problems of erosion and sedimentation, and so on.

The third factor characterizing the runs is the run-intensity, which is defined throughout this study as the average intensity or the ratio of run-sum to run-length. This quantity of run-intensity may be used as an index for the classification of regions with respect to precipitation patterns. In this study, the probability distributions of these three quantities will be obtained taking into account several possible cases of the original variable, which is the amount of precipitation in a unit of time.

First, since the unit of time for the precipitation measurement is one year, three different situations are then considered:

- (a) One single process of annual precipitation.
- (b) Two processes of annual precipitation that are mutually independent.
- (c) Two processes of annual precipitation that are dependent.

The term "process" is used in the narrow sense of "stochastic process." It is assumed that any functional dependence on time, such as trend or periodicity, has been removed from any process under consideration. The total amount of annual precipitation is considered as the original random variable.

*Former Ph.D. Graduate of Colorado State University, Civil Engineering Department, Fort Collins, Colorado, presently Adjoint Director of Hydraulic Management, Ministere des Richesses Naturelles, Quebec, P.Q. Canada.

**Professor, Department of Mathematics and Statistics, Colorado State University, Fort Collins, Colorado.

With respect to the original process, two alternative assumptions are made:

- (a) The annual precipitations are independent identically distributed normal random variables.
- (b) The annual precipitations are independent identically distributed gamma-type random variables.

The hypothesis of independence in annual precipitation is supported by Markovic [3], and physically speaking, seems to be realistic because only some factors of small effects (carry-over of water in river basins, evaporation, etc.) may affect the amount of precipitation of the following year. This hypothesis may be easily verified (autocorrelation test, run test, etc.) before analyzing data for positive and negative runs.

The hypothesis of normality of the annual precipitation is, in fact, one of truncated normality (no negative precipitation) at origin. In the regions where the probability of zero annual precipitation is high (arid or arctic regions), neither the hypothesis of normality nor the hypothesis of gamma distribution are applicable.

From the analysis of the samples from 1141 stations in the western United States, Markovic [3] found that, on the average, the annual precipitations are positively skewed, and the gamma distribution hypothesis is more realistic than the hypothesis of normality.

Second, the three main variables, run-length, run-sum, and run-intensity are analyzed also with monthly precipitation as the random variables. In this case the hypothesis of independence is tested and stated at the beginning of analysis, and no hypothesis of distribution of monthly precipitation is made.

The critical level (water demand) considered in this study is the mean value of the process. In the case of annual precipitation, the second order stationarity of the process is assumed. Therefore, the critical level is assumed to be invariant in time. In the case of monthly precipitation, the stationarity is obtained by standardization of the process.

In order to simplify the algebraic operations, the annual precipitation series is standardized, and in both cases (annual and monthly precipitation) the critical level is assumed to be zero.

Chapter II

A SEQUENCE OF INDEPENDENT VARIABLES

2.1 Introduction. In this chapter, a single process of annual precipitation is analyzed in order to obtain the statistical properties of the three main variables characterizing the positive and negative runs: run-length, run-sum, and run-intensity. This type of single record analysis is necessary before one can study several series and obtain correlation properties of one station with another or of one region with another.

2.2 Formulation of the problem. The problem was formulated by Downer, Siddiqui and Yevjevich [1]. For the sake of ready reference, however, it seems desirable to summarize the essentials of that paper. Some of their results are reported here in a strengthened form, and some new results are also included.

Let X_n , $n = 1, 2, \dots$, be independent identically distributed random variables with a common distribution F , which is assumed to be continuous. In the application to be followed after the derivation of theoretical results, X_n is the total precipitation at a given station during the n -th year. However, it can also represent the sum of precipitations over several stations in a given region. Also the unit of time may be shorter or longer than a year.

A level, c , in the range of values of X_n is chosen such that $0 < F(c) < 1$, and the n -th year is classified as a surplus year if $X_n > c$ and, in that case, refer to $X_n - c$ as the surplus. Similarly, the n -th year is called a deficit year if $X_n \leq c$, in which case, $c - X_n$ is called the deficit. Thus defined, these surpluses and deficits are all positive random variables.

A consecutive sequence of k surplus years preceded and succeeded by a deficit year is called a positive run-length k , the sum of surpluses $\sum(X_n - c)$ over such a run is the positive run-sum, and this run-sum divided by the run-length is called the positive run-intensity. Similar definitions hold for negative runs.

For $j = 1, 2, \dots$, let N_{1j} denote the length of the j -th negative run-length and N_{2j} the length of the following positive run-length. If the initial observation, X_1 , is greater than c , the initial positive run is disregarded. Suppose that the j -th negative run starts with X_{i+1} . Set

$$\begin{aligned} S_{1j} &= \sum_{k=1}^{N_{1j}} (c - X_{i+k}), & I_{1j} &= \frac{S_{1j}}{N_{1j}}, \\ S_{2j} &= \sum_{k=1}^{N_{2j}} (X_{i+N_{1j}+k} - c), & I_{2j} &= \frac{S_{2j}}{N_{2j}} \end{aligned} \quad (2.1)$$

Then S_{1j} , S_{2j} are the negative and positive run-sums,

respectively, for the j -th run, and I_{1j} , I_{2j} are the corresponding run-intensities. The properties of N_{ij} , S_{ij} , I_{ij} , $i = 1, 2$, $j = 1, 2, \dots$ are studied in the further text.

2.3 The independence of $\{(N_{1j}, S_{1j})\}$ and $\{(N_{2j}, S_{2j})\}$. For convenience the following notations are introduced:

$$\begin{aligned} p &= F(c) = P(X_n \leq c), & q &= 1 - p; \\ F_1(x) &= \frac{F(c) - F(c-x)}{F(c)}, & \text{if } x \geq 0, \\ &= 0 & \text{if } x < 0; \\ F_2(x) &= \frac{F(x+c) - F(c)}{1 - F(c)}, & \text{if } x \geq 0, \\ &= 0 & \text{if } x < 0. \end{aligned} \quad (2.2)$$

Let X_{1n}^* , $n = 1, 2, \dots$, be a sequence of independent random variables each with the distribution F_1 and X_{2n}^* , $n = 1, 2, \dots$, another sequence of independent random variables, independent of the sequence X_{1n}^* , with the distribution F_2 . Then

$$\begin{aligned} P(X_{1n}^* \leq x) &= P(c - X_n \leq x | X_n \leq c) = F_1(x), \\ P(X_{2n}^* \leq x) &= P(X_n - c \leq x | X_n > c) = F_2(x), \\ P(\sum_{j=1}^m X_{1j}^* \leq x) &= P(\sum_{j=1}^m (c - X_j) \leq x | X_j \leq c, j = 1, \dots, m) = F_1^{\otimes m}(x), \\ P(\sum_{j=1}^m X_{2j}^* \leq x) &= P(\sum_{j=1}^m (X_j - c) \leq x | X_j > c, j = 1, \dots, m) = F_2^{\otimes m}(x) \end{aligned} \quad (2.3)$$

where, for any distribution function H and $m = 1, 2, \dots, H^{\otimes m}$ denotes the m -fold convolution of H with itself.

First consider the distribution of N_{1j} . If $X_1 \leq c$, then $P(N_{11} = k | X_1 \leq c) = P(X_i \leq c, i = 1, \dots, k, X_{i+k} > c | X_1 \leq c) = qp^{k-1}$, $k = 1, 2, \dots$.

If $X_1 > c$, then

$$\begin{aligned} P(N_{11} = k | X_1 > c) &= \sum_{j=1}^{\infty} P(X_i > c, i = 1, \dots, j, X_{j+i} \leq c, \\ &i = 1, \dots, k, X_{j+k+1} > c | X_1 > c) = p^k \sum_{j=1}^{\infty} q^j = qp^{k-1}, \\ &k = 1, 2, \dots \end{aligned}$$

Hence, the unconditional distribution of N_{11} is

$$\begin{aligned} P(N_{11} = k) &= P(N_{11} = k | X_1 \leq c) P(X_1 \leq c) + P(N_{11} \\ &= k | X_1 > c) P(X_1 > c) = qp^{k-1}, \quad k = 1, 2, \dots . \end{aligned} \quad (2.4)$$

Similarly,

$$P(N_{11} = k_1, N_{21} = k_2) = p^{k_1} q^{k_2}; \quad P(N_{21} = k_2) p q^{k_2-1},$$

so that

$$P(N_{11} = k_1, N_{21} = k_2) = P(N_{11} = k_1) P(N_{21} = k_2),$$

and N_{11} and N_{21} are independent. This argument can be extended to show that $N_{11}, N_{21}, N_{12}, N_{22}, \dots$ are mutually independent, $\{N_{1j}\}$ are identically distributed, and $\{N_{2j}\}$ are identically distributed.

Now, look at the joint distribution of (N_{11}, S_{11}) . From (2.3) it follows that

$$P(S_{11} \leq x | N_{11} = k) = F_1^{\otimes k}(x), \quad (2.5)$$

hence

$$P(N_{11} = k, S_{11} \leq x) = qp^{k-1} F_1^{\otimes k}(x), \quad (2.6)$$

Similar expressions hold for (N_{21}, S_{21}) . Finally,

$$F_S(x) = P(S_{11} \leq x) = \sum_{k=1}^{\infty} qp^{k-1} F_1^{\otimes k}(x). \quad (2.7)$$

Again, one can show that the sequence of vectors (N_{1j}, S_{1j}) is mutually independent and identically distributed with (2.6). This sequence is also independent of (N_{2j}, S_{2j}) , which themselves are mutually independent and identically distributed. Since the treatment of one vector sequence is exactly parallel to the other, only one is considered. (In fact $X_n \leq c$ is equivalent to $-X_n > -c$ so that a negative run for X_n at level, c , is equivalent to a positive run for $-S_n$ at level, $-c$). We choose to concentrate on the negative run (N_{1j}, S_{1j}) . We drop the subscript, j , and write it as (N_1, S_1) unless the whole sequence is considered.

2.4 The distribution of S_1 in some special cases.

From (2.7), the distribution function, F_S , of S_1 , is directly related to F_1 rather than to F . Since $0 < p < 1$, $p^n \rightarrow 0$, terms after some $k = n$ may be negligible. For example, if $p = 1/2$, $p^7 < 0.01$ and the series may be truncated at the sixth term with the error of approximation less than one percent uniformly for all x . Actually, since $F_1^{\otimes k}(x) \leq 1$, then

$$F_S(x) = \sum_{k=1}^n qp^{k-1} F_1^{\otimes k}(x) \leq \sum_{k=n+1}^{\infty} qp^{k-1} = p^n \quad (2.8)$$

For example, let $F_1(x) = F(x, \lambda, r)$ with the density

function

$$f(x, \lambda, r) = \frac{\lambda^r x^{r-1} e^{-\lambda x}}{\Gamma(r)}, \quad x > 0,$$

and $r \geq 0$. If $r = 1$, then

$$\begin{aligned} f_S(x) &= F'_S(x) = q\lambda e^{-\lambda x} [1 + p\lambda x + \frac{(p\lambda x)^2}{2!} + \dots] \\ &= q\lambda e^{-\lambda x}, \quad x > 0, \end{aligned} \quad (2.9)$$

an exponential distribution. For arbitrary $r > 0$,

$$\begin{aligned} F_S(x) &= q[F(x, \lambda, r) + p F(x, \lambda, 2r) + p^2 F(x, \lambda, 3r) + \dots] \\ &= q \sum_{k=1}^n p^{k-1} F(x, \lambda, kr) + R_n(x), \end{aligned}$$

where

$$\begin{aligned} R_n(x) &= q[p^n F(x, \lambda, (n+1)r) + p^{n+1} F(x, \lambda, (n+2)r) + \dots] \\ &\leq qp^n F(x, \lambda, (n+1)r)[1 + p + p^2 + \dots] = p^n F(x, \lambda, (n+1)r). \end{aligned} \quad (2.10)$$

This follows because, for $x > 0$,

$$F(x, \lambda, kr) < F(x, \lambda, (k-1)r).$$

Usually, $n = 2$ or 3 may give a satisfactory approximation.

2.5 Moment generating functions. If y is any random variable with the distribution function, $J(y)$, and E is the mathematical expectation operator, then

$$M_y(\theta) = E e^{\theta y} = \int_{-\infty}^{\infty} e^{\theta y} dJ(y)$$

is called the moment generating function of Y or of the distribution, J . θ is taken to be a complex number and $M_y(\theta)$ exists at least for $\operatorname{Re} \theta \leq 0$. If for some $r = 1, 2, \dots$, the r -th moment of Y exists, then it is given by

$$\mu_r'(y) = E Y^r = M_y^{(r)}(0),$$

where $M_y^{(r)}(0)$ is the r -th derivative of M evaluated at $\theta = 0$.

If Y_1, Y_2, \dots, Y_n are independent and $Y = Y_1 + \dots + Y_n$, then

$$M_y(\theta) = M_{Y_1}(\theta) M_{Y_2}(\theta) \dots M_{Y_n}(\theta).$$

The function, $K_y(\theta) = \ln M_y(\theta)$, is called the cumulant generating function of Y . The r -th cumulant of Y

exists if the r -th moment of Y exists and is given by

$$k_r(y) = r\text{-th cumulant of } Y = K_y^{(r)}(0).$$

Clearly, for $Y = Y_1 + \dots + Y_n$, where Y_1, \dots, Y_n are independent

$$K_y(\theta) = K_{Y_1}(\theta) + \dots + K_{Y_n}(\theta)$$

$$k_r(y) = k_r(y_1) + \dots + k_r(y_n).$$

Recall that $\{X_{ln}\}$ have the common distribution function, F_1 . Set

$$M_1(v) = Ee^{\theta X_{11}^*} = \int_{-\infty}^{\infty} e^{\theta x} dF_1(x) = \frac{e^{\theta v}}{p} \int_{-\infty}^c e^{-\theta x} dF(x),$$

$$K_1(v) = \ln M_1(v).$$

From Downer, Siddiqui and Yevjevich [1],

$$M_1(u, v) = Ee^{uN_1 + vS_1} = \frac{q \exp[u - K_1(v)]}{1-p \exp[u - K_1(v)]},$$

$$K_1(u, v) = \ln M_1(u, v).$$

Also

$$EN_1 = \frac{1}{q}, \quad \text{var } N_1 = \frac{p}{q^2}$$

$$ES_1 = \frac{EX_{11}^*}{q}, \quad \text{var } S_1 = \frac{q \text{ var } X_{11}^* + p(EX_{11}^*)^2}{q^2}$$

$$\text{Cov}(N_1, S_1) = \frac{p}{q^2} = EX_{11}^*$$

$$\rho(N_1, S_1) = \frac{pEX_{11}^*}{\sqrt{pq \text{ var } X_{11}^* + p^2(EX_{11}^*)^2}} \quad (2.12)$$

where $\text{var } X$ is the variance of X , and $\text{cov } (X, Y)$ is the covariance, and $\rho(X, Y)$ is the correlation between X and Y .

The authors just mentioned did not give the moment generating function of $I_1 = \frac{S_1}{N_1}$ but, in a similar argument,

$$M_{I_1}(\theta) = Ee^{\theta I_1} = Ee^{\theta \frac{S_1}{N_1}} = \sum_{n=1}^{\infty} qp^{n-1} \left\{ M_1\left(\frac{\theta}{n}\right) \right\}^n. \quad (2.13)$$

The evaluation of the moments of I_1 involves sums of the form

$$A_r(z) = \sum_{n=1}^{\infty} \frac{z^{n-1}}{n^r}, \quad r = 1, 2, \dots, 0 < z < 1.$$

Now,

$$A_1(z) = 1 + \frac{z}{2} + \frac{z}{3} + \dots = -\frac{1}{z} \ln(1-z).$$

Integrating both sides from 0 to p gives

$$-\int \frac{p}{z} \ln(1-z) dz = p + \frac{p^2}{2^2} + \frac{p^3}{3^2} + \dots + \frac{p^n}{n^2} + \dots = \frac{1}{p} A_2(p),$$

and so on.

Finally,

$$EI_1 = EX_{11}^*, \quad \text{var } I_1 = \frac{q}{p} \text{ var } X_{11}^* \ln\left(\frac{1}{q}\right). \quad (2.14)$$

There is little point in giving the algebraic form of higher moments as they can be numerically calculated in a specific situation.

Chapter III

NORMAL AND GAMMA DISTRIBUTED VARIABLES

3.1 Normal sequence. Suppose that the original sequence $\{\bar{X}_n\}$ is an independent normally distributed sequence with $E\bar{X}_n = 0$, $\text{var } \bar{X}_n = 1$. (If $E\bar{X}_n = \mu$ and $\bar{X}_n = \sigma^2$, then consider the standardized sequence $(\bar{X}_n - \mu)/\sigma$) Downton, Siddiqui and Yevjevich [1] have studied this situation exhaustively for (N_{2j}, S_{2j}) . Their results are equally applicable to (N_{1j}, S_{1j}) . Now consider the case of $c = E\bar{X}_n = 0$ to illustrate a method of approximating to the distribution of S_{11} and I_{11} . Thus

$$ES_1 = 1.59577, \text{ var } S_1 = 2.0, ES_1^3 = 18.8615, ES_1^4 = 93.0225$$

where an approximation to the value of π is used. Also

$$EI_1 = 0.797885, EI_1^2 = 0.888496, EI_1^3 = 1.263186,$$

$$EI_1^4 = 2.064387, EN_1 = 2, \text{ var } N_1 = 2, \text{ cov}(N_1, S_1) =$$

$$= 1.59577, \rho(N_1, S_1) = 0.798.$$

The coefficients of skewness of S_1 is

$$C(S_1) = 1.82.$$

Since this is positive, a gamma distribution is chosen to approximate to $f_s(x)$. Siddiqui [4] gives the method for this type of approximation.

Let

$$f(x; g, h) = \frac{e^{-x/2g} x^{k/2-1}}{(2g)^{h/2} \Gamma(h/2)} \quad \text{for } x > 0 \\ = 0 \quad \text{otherwise,} \quad (3.1)$$

be the probability density function of a gamma variate, to which the probability density function of S_1 is approximated.

In (3.1), g is a scale factor and h is the effective number of degrees of freedom.

The approximation to the probability density of S_1 can be improved as follows:

$$f(x) = f(x; g, h) \sum_{m=0}^{\infty} \frac{m! \Gamma(h/2)}{\Gamma(m+h/2)} \frac{d_m}{(2g)^m} L_m^{(\frac{h}{2}-1)} \left(\frac{x}{2g}\right) \quad (3.2)$$

where $L_m^{(c)}(y)$ is the Laguerre polynomial of degree, m .

$$L_m^{(c)}(y) = \sum_{j=0}^m \binom{m+c}{m-j} \frac{(-y)^j}{j!} \quad (3.3)$$

and

$$d_m = \sum_{j=0}^m \binom{m-1+h/2}{m-j} (-1)^j (2g)^{m-j} \frac{\gamma_j}{j!} \quad (3.4)$$

where

$$\gamma_j = E(S_1)^j.$$

The parameters g and h can be computed by the method of moments, i.e., equating the first two moments of the probability density function in Eq. (3.1) with the moments, ES_1 and $E(S_1)^2$, already found. The first two moments of the distribution in Eq. (3.1) are gh and $g^2h(h+2)$. Thus, setting $gh = ES_1$ and $g^2h(h+2) = ES_1^2$,

$$g = 0.626657, h = 2.546482.$$

Then

$$f(x; g, h) = 0.816398 e^{-0.797885x} x^{0.273241}$$

In this kind of approximation, only the first few polynomials are really important. As a general rule, the order, m , of the last polynomial considered must be such that:

- (a) No appreciable oscillations appear in the probability density function.
- (b) The coefficient of x^m must be small in comparison with the coefficients of the terms of lower order.

With those considerations, the probability density function of S_1 is truncated at $m = 4$.

Table 1 shows the different computations. In this table, $A_m = \frac{m! \Gamma(h/2)}{\Gamma(m+h/2)}$.

Finally, the probability density function of S_1 (and S_2) is

$$f_s(x) = 0.816398 e^{-0.797885x} x^{0.273241} (0.790207 + \\ + 0.514732x - 0.265132x^2 + 0.042132x^3 - 0.001922x^4).$$

3.2 Approximated probability density functions of I_1 and I_2 . As before, the probability density function of I_1 (and I_2) will be approximated by a function of a gamma-variate.

In this case,

$$g = \frac{\text{Var } I_1}{2E I_1} \approx 0.157840$$

TABLE 1
IMPROVEMENT OF PROBABILITY DENSITY FUNCTION OF S_i ($i = 1, 2$)

m	$\frac{d_m}{(2g)^m}$	A_m	L_m	(0.273241) $(\frac{x}{2g})$
0	1	1		1
1	0	-		-
2	0	-		-
3	-0.017781	0.633311	1.579002-2.968478x+1.041905x ² -0.084658x ³	
4	-0.192013	0.592816	1.686864-4.228341x+2.226157x ² -0.361764x ³ + +0.016887x ⁴	

$$h = \frac{2(EI_1)^2}{\text{Var } I_1} = 5.055031 .$$

The parameters d_m and A_m and the functions L_m , for $m = 0$ to 4, are given in Table 2.

Finally, the probability density function of I_1 and I_2 is

$$f_I(x) = 13.650570e^{-3.167764x} x^{1.527516} (0.595968 + 2.070133x - 2.848625x^2 + 1.356767x^3 - 0.198404x^4) . \quad (3.5)$$

Figures 1 and 2 show the probability density functions of S_1 (or S_2) and I_1 (or I_2).

3.3 Gamma distributed sequence. Let a random variable, X , have the distribution function, $F(x)$, with the probability density function

$$f(x) = \frac{x^{r-1} e^{-x}}{\Gamma(r)} , \quad \text{if } x > 0 \\ = 0 , \quad \text{if } x \leq 0 ,$$

where $r > 0$. One can introduce a scale factor λ , but it will simply involve multiplying the k -th moment by λ^k . Since $EX = r$, $\text{var } X = r$, we consider the moments of, the standardized variable

$$X_1 = \frac{x - r}{\sqrt{r}} ,$$

and the sequence X_n , $n = 1, 2, \dots$, which are identically distributed.

If X_{11}^* and X_{12}^* denote the truncated random variables with $c = 0$, then

$$EX_{11}^{*k} = r^{-k/2} EX_{11}^k ,$$

where X_{11} is the variable, X , truncated at $EX = r$.

Similarly, $EX_{21}^{*k} = r^{-k/2} EX_{22}^k$.

Let F_1 and F_2 be the distribution functions obtained from Eq. (2.2). Then

TABLE 2
IMPROVEMENT OF PROBABILITY DENSITY FUNCTION OF I_i ($i = 1, 2$)

m	$\frac{d_m}{(2g)^m}$	A_m	L_m	(1.527516) $(\frac{x}{2g})$
0	1	1		1
1	0	-		1
2	0	-		-
3	0.035602	0.148638	6.727776-25.295999x+22.716197x ² -5.297942x ³	
4	-0.439633	0.107562	9.296972-46.608005x+62.782081x ² -29.284471x ³ + +4.195661x ⁴	

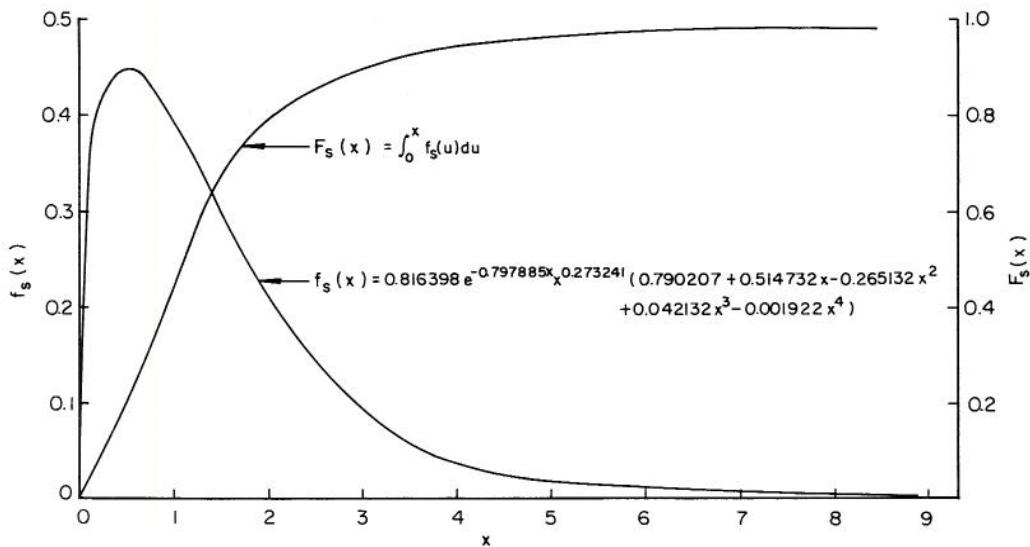


Figure 1 Distribution function and probability density function of S_1 and S_2

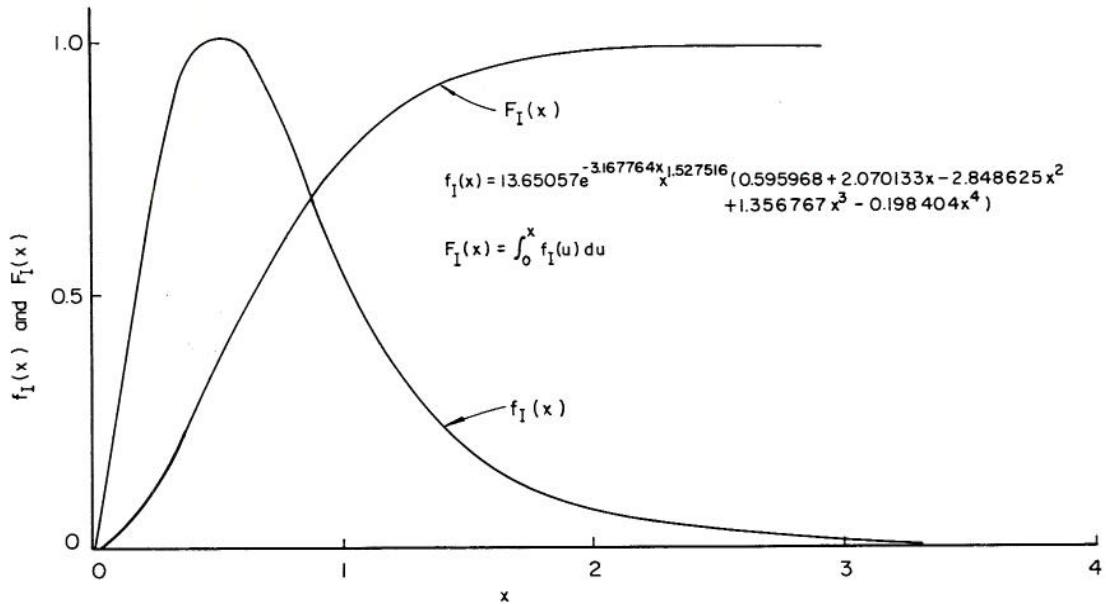


Figure 2 Distribution function and probability density function of I_1 and I_2

$$dF_1(x) = \frac{(r-x)^{r-1} e^{-(r-s)}}{\Gamma(r) P(r,r)} dx, \quad 0 \leq x \leq r$$

where

$$P(a,x) = \frac{1}{\Gamma(a)} \int_0^x t^{a-1} e^{-t} dt \quad a > 0, x > 0,$$

is the incomplete gamma function.

Then

$$\begin{aligned} \Gamma(r) P(r,r) E(X_{11}^k) &= \int_0^r x^k (r-x)^{r-1} e^{-(r-x)} dx \\ &= \int_0^r y^{r-1} (r-y)^k e^{-y} dy \end{aligned}$$

$$= \sum_{j=0}^k (-1)^j r^{k-j} \Gamma(r+j) P(r+j, r).$$

Hence,

$$E(X_{11}^*)^k = \frac{\sum_{j=0}^k (-1)^j r^{k-j} \Gamma(r+j) P(r+j, r)}{r^{k/2} \Gamma(r) P(r, r)}. \quad (3.6)$$

In a similar fashion

$$\begin{aligned} E(X_{21}^*)^k &= \frac{\sum_{j=0}^k (-1)^j \binom{k}{j} r^j \Gamma(r+k-j)}{r^{k/2} \Gamma(r) P(r, r)} - \\ &- (-1)^k E(X_{11}^*)^k \end{aligned} \quad (3.7)$$

In the distribution of N_{11} and N_{21} ,

$$p = F(r) = P(r, r), q = 1-p.$$

We define S_1 and S_2 in terms of the "normalized" variables X_{1n}^* , X_{2n}^* and then calculate their moments.

The following table shows the values of the first four moments of X_{11}^* and X_{21}^* for several values of r .

The values of the Incomplete Gamma Function have been taken from K. Pearson [5].

TABLE 3
MOMENTS OF X_{ii}^* ($i=1,2$) FOR SEVERAL VALUES OF r

r	$E(X_{11}^*)$	$E(X_{11}^*)^2$	$E(X_{11}^*)^3$	$E(X_{11}^*)^4$	$E(X_{21}^*)$	$E(X_{21}^*)^2$	$E(X_{21}^*)^3$	$E(X_{21}^*)^4$
1	0.58198	0.41802	0.32788	0.27027	0.58198	1.16396	3.49184	13.96753
2	0.64412	0.54456	0.51807	0.53441	0.64412	1.13809	2.89771	9.56150
4	0.69031	0.65486	0.72614	0.87779	0.69031	1.11027	2.49127	7.06530
6	0.71055	0.71012	0.84056	1.10099	0.71055	1.09454	2.31406	6.11766

The following tables show the first four moments of S_i and I_i for several values of r .

TABLE 4
MOMENTS OF S_i ($i=1,2$) FOR SEVERAL VALUES OF r

r	ES_1	$E(S_1)^2$	$E(S_1)^3$	$E(S_1)^4$	ES_2	$E(S_2)^2$	$E(S_2)^3$	$E(S_2)^4$
1	1.58198	4.30027	17.20120	91.63671	0.92068	2.46503	9.89973	53.01123
2	1.58768	4.33840	17.35705	92.41097	1.08383	2.86815	11.18544	57.70780
4	1.59252	4.38432	17.63093	94.21243	1.21849	3.24693	12.64762	65.03209
6	1.59359	4.40703	17.77899	95.26155	1.28230	3.44160	13.46745	69.57820

TABLE 5
MOMENTS OF I_i ($i=1,2$) FOR SEVERAL VALUES OF r

r	ES_1	$E(S_1)^2$	$E(S_1)^3$	$E(S_1)^4$	ES_2	$E(S_2)^2$	$E(S_2)^3$	$E(S_2)^4$
1	0.58198	0.38486	0.27422	0.20031	0.58198	0.98911	2.63477	9.77068
2	0.64412	0.49475	0.42168	0.37902	0.64412	0.96617	2.15717	6.50659
4	0.69031	0.59059	0.57961	0.60859	0.69031	0.94718	1.85078	4.74713
6	0.71055	0.63827	0.66659	0.75612	0.71055	0.93751	1.72266	4.09933

Comparing the moments of S_i and I_i , $i = 1,2$, obtained in this way with the same moments as for the normal, it follows that the moments corresponding to the gamma distribution of the original random variable, X_1 , converge to the moments corresponding to the normal distribution of X_1 . This convergence is almost independent of r for the moments of S_1 , but for the other random variables, S_2 , I_1 and I_2 , both assumptions are

similar for large values of r only as shown in Figs. 3, 4 and 5.

3.4 Example: Fort Collins, Station No. 5.3005.

Years of records: $N = 69$

Mean: $\mu = 14.62$

Standard deviation: $\sigma = 4.00$

Equating the mean and variance, it follows

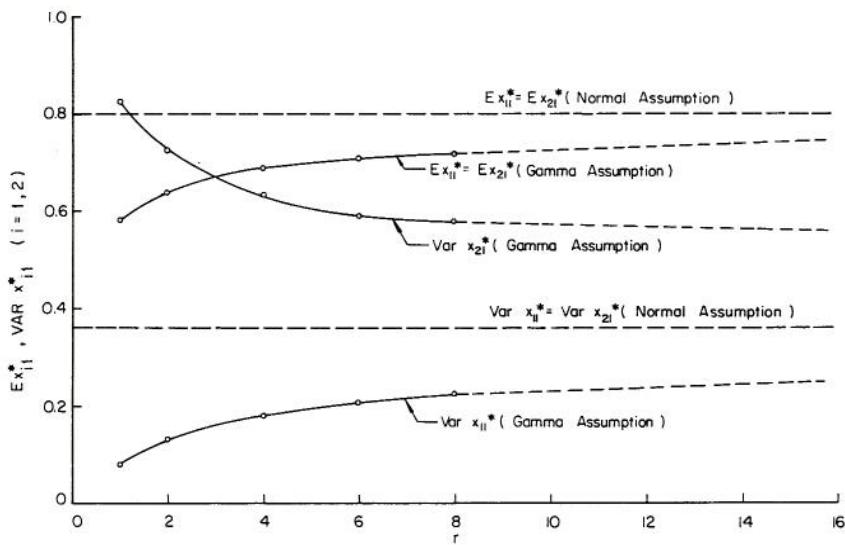


Figure 3 Expected values and variances of X_{11}^* and X_{21}^* for normal and gamma assumptions

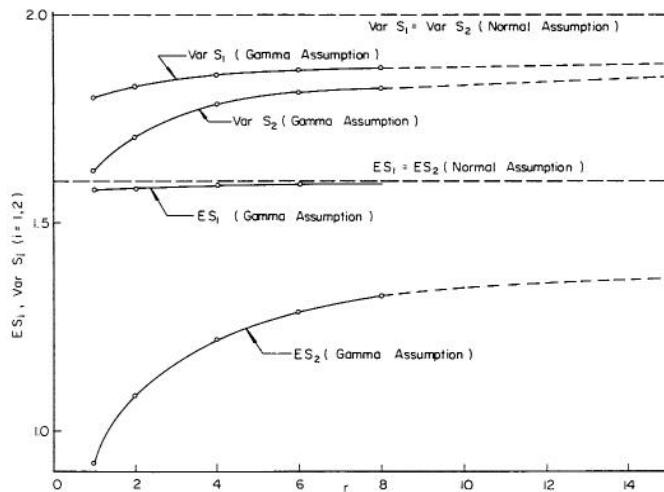


Figure 4 Expected values and variances of S_1 and S_2 for normal and gamma assumptions

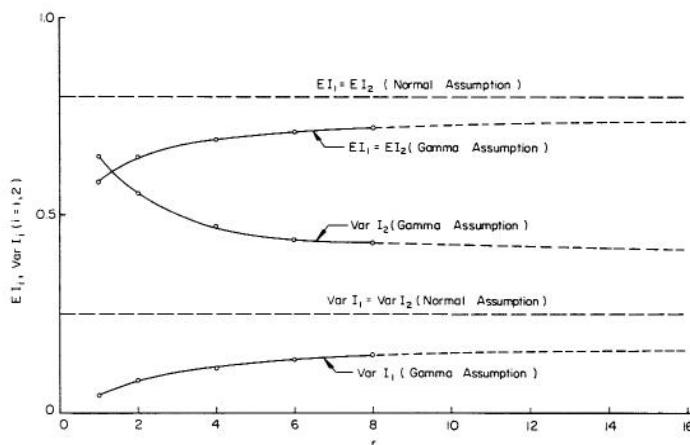


Figure 5 Expected values and variances of I_1 and I_2 for normal and gamma assumptions

TABLE 6
EXPECTED VALUES AND VARIANCES OF X_i and N_i ($i=1,2$)
FOR DIFFERENT HYPOTHESES OF $\{X_n\}$

Hypothesis of $\{X_n\}$	EX_{11}^*	EX_{21}^*	$VarX_{11}^*$	$VarX_{21}^*$	EN_1	EN_2	$VarN_1$	$VarN_2$
Normal	0.79789	0.79789	0.36338	0.36338	2.00000	2.00000	2.00000	2.00000
Gamma	0.740	0.740	0.245	0.565	2.166	1.857	2.527	1.760
From the data	0.670	0.910	0.243	0.502	1.850	1.524	0.928	0.725

TABLE 7
EXPECTED VALUES AND VARIANCES OF S_i AND I_i ($i=1,2$)
FOR DIFFERENT HYPOTHESES OF $\{X_n\}$

Hypothesis of $\{X_n\}$	ES_1	ES_2	$VarS_1$	$VarS_2$	EI_1	EI_2	$VarI_1$	$VarI_2$
Normal	1.59577	1.59577	2.00000	2.00000	0.79789	0.79789	0.25188	0.25188
Gamma	1.595	1.360	1.880	1.845	0.73800	0.73800	0.160	0.420
From the data	1.239	1.386	0.816	1.027	0.673	0.957	0.123	0.586

$$\mu = \frac{r}{\lambda} = 14.62$$

$$\sigma^2 = \frac{r}{\lambda^2} = 16.$$

Then

$$\lambda = 0.9138 \quad r = 13.360$$

The preceding example shows that the first moment of all random variables, obtained from the data, agrees quite well with the first moment of the theoretical hypothesis (better if the comparison is done with gamma hypothesis). For the random variables, N_i, S_i , and I_i ($i=1,2$), the disagreement between the higher moments in both cases, provided by the fact that the sample size and consequently the number of runs, is very small in this example; therefore, the estimation is obviously subject to large sampling fluctuations.

Chapter IV

TWO MUTUALLY INDEPENDENT PROCESSES

4.1 Introduction. In previous chapters, the parameters defining the negative and positive runs of annual precipitation were studied considering one single sequence of original random variables: the total amount of annual precipitation at one station. The concept of runs defined in this way can be generalized to several points in space simultaneously in order to study the behavior of those phenomena in the joint dimensions of time and space. This situation is often encountered in hydrology. For example, if a river is passing through two regions with similar or different meteorological conditions, the expected runs in a downstream storage project depend on the combined pattern of precipitation in both regions. In this case, two different sequences will be required in order to define the process. The same problem can arise in a large watershed in regard to the particular model of precipitation on its main tributaries.

4.2 Formulation of the problem. Consider a sequence of a two-dimensional process, (X_n, Y_n) , $n = 1, 2, \dots$, where these vectors are mutually independent and have a common distribution function, $F(x, y)$. Given two levels, c_1 and c_2 , such that $0 < F(c_1, c_2) < 1$, we have four possible events:

$$A_n = \{X_n \leq c_1, Y_n \leq c_2\} \quad B_n = \{X_n \leq c_1, Y_n > c_2\}$$

$$C_n = \{X_n > c_1, Y_n \leq c_2\} \quad D_n = \{X_n > c_1, Y_n > c_2\}$$

Of these four, A_n and D_n are of interest to us.

The n -th year will be called deficit for both sequences if A_n occurs and surplus if D_n occurs. A sequence of k consecutive A 's followed and preceded by any other event is a negative run of length, k . A sequence of k consecutive D 's followed and preceded by any other event is a positive run of length, k . (For the initial run the requirement of "preceded by" is dropped.) The situation is depicted in Fig. 6.

$$P(A_n) = F(c_1, c_2) = p, \text{ say, } P(B_n \cup C_n \cup D_n) = P(A_n^c) = 1 - p = q.$$

Thus the distribution of N_{11} is still given by the formula

$$P(N_{11} = k) = qp^{k-1}, \quad k = 1, 2, \dots$$

The difference is that now there is no guarantee that a negative run will be immediately followed by a positive run. In fact, it is quite possible that a negative run is followed by a few B and C type events, which in turn, are followed by another negative run. Also here $q \neq P(D_n)$. Since the discussion of the positive runs is parallel to that of negative runs, we omit their mention entirely. We now use the symbols S_{11} for $\sum(c-X_n)$ and S_{21} for $\sum(c-Y_n)$, where the summation is over a (common) negative run.

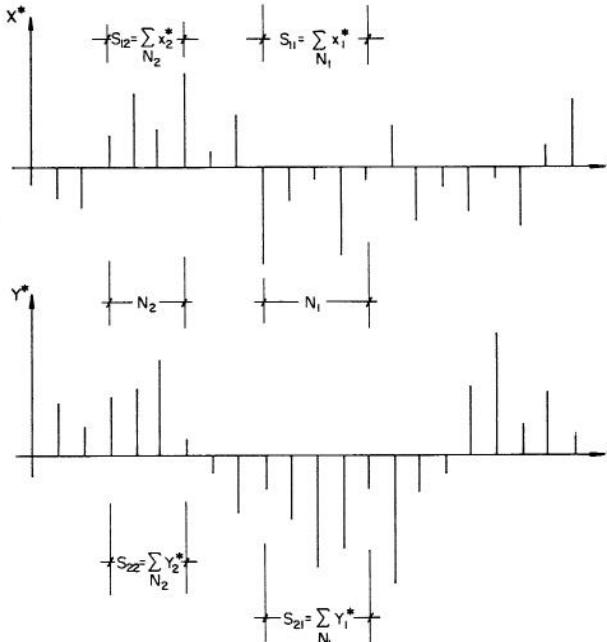


Figure 6 Graphical representation of the random variables S_{11} , S_{12} , S_{22} , N_1 and N_2

When $\{X_n\}$ is independent of $\{Y_n\}$, we have

$$F(c_1, c_2) = F(c_1)G(c_2),$$

where $F(x) = P(X_n \leq x)$, $G(y) = P(Y_n \leq y)$. For example, if both X_n and Y_n are standard normal and $c_1 = c_2 = 0$, then $F(0, 0) = 1/4$. However, we are at liberty to choose F and G differently, for instance, F to be normal and G to be a gamma distribution.

Now, let

$$\begin{aligned} F_{12}(x, y) &= P(c_1 - X_1 \leq x, c_2 - Y_1 \leq y | X_1 \leq c_1, Y_1 \leq c_2) \\ &= P(c_1 - X_1 \leq x | X_1 \leq c_1) P(c_2 - Y_1 \leq y | Y_1 \leq c_2), \end{aligned} \quad (4.1)$$

so that the random variables, $\{X_{1n}\}$ and $\{Y_{1n}\}$, can be defined independently by the truncation of F and G , respectively. The entire discussion of Chapter II carries through for S_{1j} and S_{2j} except for their covariance properties. We have

$$\begin{aligned} E S_{11} S_{21} &= E[E(S_{11} S_{21} | N_1)] = EN_1^2 \sum_{k=1}^{N_1-1} EX_{11}^* EY_{11}^* \\ &= EN_1^2 EX_{11}^* EY_{11}^*, \end{aligned}$$

so that

$$\text{cov}(S_{11}, S_{21}) = \text{var } N_1 EX_{11}^* EY_{11}^* = \frac{p}{q^2} EX_{11}^* EY_{11}^*. \quad (4.2)$$

4.3 Comment on the dependent case. From the equation (4.1), it is apparent that no general discussion can be carried very far if $\{X_n\}$ and $\{Y_n\}$ are not independent, i.e., when $F(x,y) \neq F(x)G(y)$. The essential difficulty is in finding the joint distribution of the truncated random variables (X_{11}^*, Y_{11}^*) .

Even the marginal distribution of X_{11}^* (or of Y_{11}^*) depends on the joint condition $\{X_1 \leq c_1, Y_1 \leq c_2\}$.

Chapter V

MONTHLY PRECIPITATION SERIES

5.1 Introduction. In the three previous chapters the annual precipitation as the basic random variable leading to an objective definition of runs was considered. The series in this form were independent within and strictly stationary.

However, in some cases it is preferable to reduce the length of this original random variable in order to create another process in which the length of observations will be longer. In practical terms this new process offers more advantages, in particular, in the study of those phenomena renewable in a short period of time. For example, the drought, defined as the negative run over the mean of annual precipitation, does not mean anything to a farmer so long as the precipitation is concentrated in the right period.

In this chapter, the monthly precipitation is the basic random variable. The time series formed by the total precipitation during a month are not stationary because of the seasonal variations. Each series must be considered as a sample of 12 different populations, and some transformations should be necessary in order to bring about stationarity.

5.2 Formulation of the problem. We consider a sequence of monthly precipitation. Let P_t , $t = \tau + 12(n-1)$ be the total amount of precipitation in the τ -th month of the n -th year. Here, $\tau = 1, 2, \dots, 12$ and $n = 1, 2, \dots$. Fix τ and set

$$x_n = \frac{P_t - \mu_\tau}{\sigma_\tau}, \quad t = \tau + 12(n-1),$$

where μ_τ is the mean value and σ_τ is the standard deviation for the month τ . x_n , $n = 1, 2, \dots$, then corresponds to the standardized values of P_t for the same month of successive years. Clearly, this may be assumed to be either an independent sequence or a mildly dependent stationary sequence. What we assume concerning the dependence is the following. Let

$$\begin{aligned} x_n &= 1, & \text{if } x_n > 0, \\ &= 0, & \text{if } x_n \leq 0. \end{aligned}$$

We assume that the sequence of x_n forms a two state Markov process with stationary transition probabilities. That is,

$$P(x_n | x_{n-1}, x_{n-2}, \dots, x_1) = P(x_n | x_{n-1}) = P(x_2 | x_1). \quad (5.1)$$

Let

$$P(x_2 = 0 | x_1 = 0) = 1 - \alpha, \quad P(x_2 = 1 | x_1 = 0) = \alpha$$

$$P(x_2 = 0 | x_1 = 1) = \beta \quad P(x_2 = 1 | x_1 = 1) = 1 - \beta$$

The transition matrix of the model is

$$P = \begin{bmatrix} 0 & \alpha & 1-\alpha \\ 1-\beta & \beta & 1-\beta \end{bmatrix},$$

with equilibrium probabilities

$$\lim_{n \rightarrow \infty} P(x_n = 0) = \pi_0 = \frac{\beta}{\alpha+\beta}, \quad \lim_{n \rightarrow \infty} P(x_n = 1) = \pi_1 = \frac{\alpha}{\alpha+\beta}.$$

We further assume that the initial probability distribution is given by

$$P(x_1 = 0) = \pi_0, \quad P(x_1 = 1) = \pi_1$$

so that the chain is stationary. That is,

$$P(x_n = 0) = \pi_0, \quad P(x_n = 1) = \pi_1, \quad \text{for all } n = 1, 2, \dots$$

Now let N_{1j} be the j -th run of 0's and N_{2j} the following run of 1's. We have

$$\begin{aligned} P(N_{11} = k | x_1 = 0) &= P(x_1 = 0, \dots, x_k = 0, x_{k+1} = 1 | x_1 = 0) \\ &= (1-\alpha)^{k-1}\alpha, \end{aligned}$$

$$\begin{aligned} P(N_{11} = k | x_1 = 1) &= \sum_{j=1}^{\infty} P(x_i = 1, i=1, \dots, j, x_{j+i} = 0, \\ &\quad i = 1, \dots, k, \dots) \end{aligned}$$

$$\begin{aligned} x_{j+k+1} = 1 | x_1 = 1) &= \sum_{j=1}^{\infty} \beta^{j-1}(1-\beta)(1-\alpha)^{k-1}\alpha \\ &= (1-\alpha)^{k-1}\alpha. \end{aligned}$$

Hence, the unconditional probability

$$P(N_{11} = k) = \alpha(1-\alpha)^{k-1}, \quad k = 1, 2, \dots,$$

which is the same as Eq. (2.4) with $p = 1 - \alpha$.

Similarly,

$$P(N_{21} = k) = \beta(1-\beta)^{k-1}, \quad k = 1, 2, \dots$$

Let

$$z_n = \sum_{j=1}^n x_j, \quad Y_n = n - z_n.$$

Then Z_n is the number of surplus months out of n , and

Y_n is the number of deficit months. We know that [6] Y_n is asymptotically ($n \rightarrow \infty$) normally distributed with

$$EY_n \sim n \frac{\beta}{\alpha+\beta}, \text{ var } Y_n \sim n \frac{\alpha\beta(2-\alpha-\beta)}{(\alpha+\beta)^3}$$

(Since $Y_n + Z_n = n$, $\text{var}(Y_n + Z_n) = 0$ for all n .)

5.3 Properties of runs. Defining S_1, I_1, S_2, I_2 as before, we note that for $c = 0$, the model outlined above is equivalent to the independent sequence model except that $q = \alpha, p = 1 - \alpha$ for (N_1, S_1, I_1) and the same (p, q) will not apply for (N_2, S_2, I_2) unless $\beta = 1 - \alpha$, which is the independent case. Thus, when discussing N_1, S_1, I_1 , i.e., negative run-length, run-sum and run-intensity, we set $p = 1 - \alpha, q = \alpha$, in the formulas (2.11, 2.12 and 2.13).

5.4 Example.

Station 4.7740 San Diego WB APT

The probability density function of monthly precipitation for this station is given in Fig. 7. We obtain the following values for the parameters:

$$\alpha = 0.290$$

$$\beta = 0.683$$

$$ES_1 = 1.596; E(S_1)^2 = 4.6066$$

$$ES_2 = 1.609; E(S_2)^2 = 5.5628$$

$$EI_1 = 0.4629; E(I_1)^2 = 0.2508$$

$$EI_2 = 1.0941; E(I_2)^2 = 2.3969$$

From the data:

$$\hat{E}S_1 = 1.5953; \hat{E}(S_1)^2 = 4.2246$$

$$\hat{E}S_2 = 1.5987; \hat{E}(S_2)^2 = 5.8756$$

$$\hat{E}I_1 = 0.4629; \hat{E}(I_1)^2 = 0.2416$$

$$\hat{E}I_2 = 1.1023; \hat{E}(I_2)^2 = 2.2147$$

5.5 Explanation of appendices. In Appendix I the following tables are provided:

1. Table of incomplete gamma function $P(a, x)$ for $a = 1(1)14$, and $x = 1, 2, 4, 6, 10$.
2. Data used in example of Chapter III.
3. Locations of precipitation stations in Colorado.
4. Table giving numerical values of means and variances of variables related to runs for the annual precipitation series at stations in Colorado.

Appendix II provides numerical values of parameters discussed in Chapter V, such as EX_1^* , α , β , π_0 , π_1 , EN_1 , for monthly precipitation series at stations in the Western United States. Areal distribution of these stations is also provided.

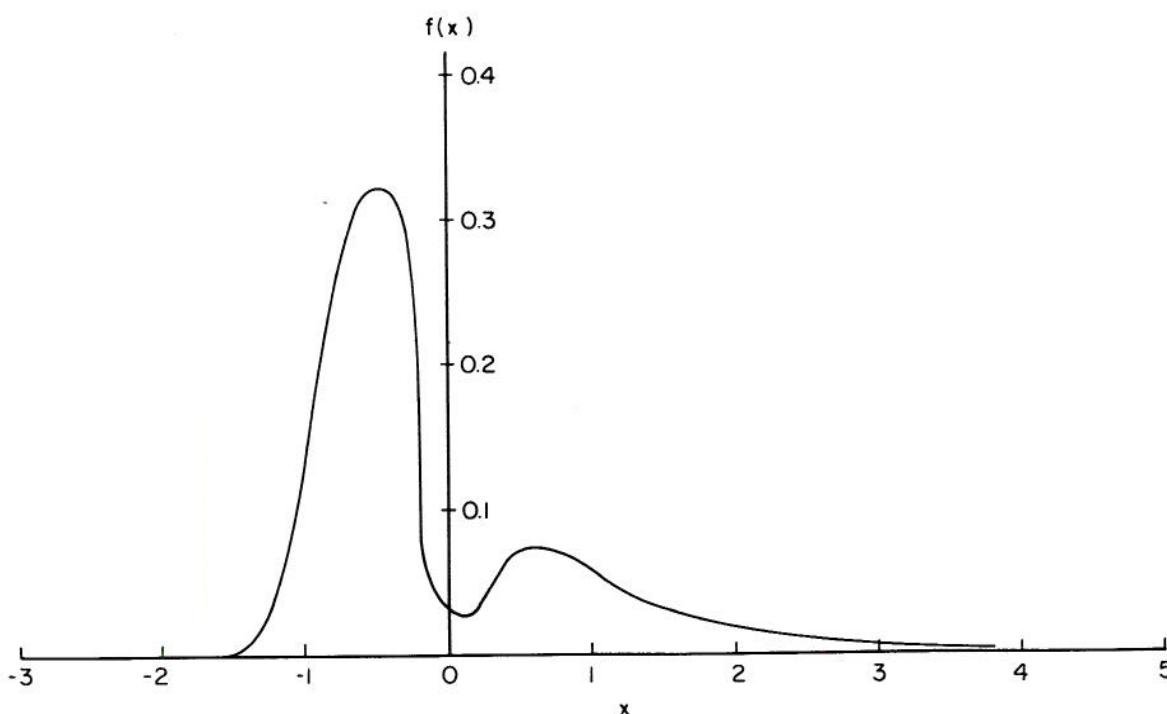


Figure 7 Probability density function of monthly precipitation. Station No. 4.7740.
San Diego W.B. APT

REFERENCES

1. Downer, R. N., Siddiqui, M. M., and Yevjevich, V., "Application of Runs to Hydrologic Droughts." Proceedings of the International Hydrology Symposium, Fort Collins, Colorado, September 1967, Vol. 1, Paper No. 63.
2. Yevjevich, V., "An Objective Approach to Definitions and Investigations of Continental Hydrologic Droughts." Colorado State University Hydrology Paper No. 23, August 1967, Fort Collins, Colorado.
3. Markovic, R. D., "Probability Functions of Best Fit to Distributions of Annual Precipitation and Runoff." Colorado State University
4. Siddiqui, M.M., "Test for Regression Coefficients When Errors are Correlated." *An. Math. Statistics*, Vol. 31, 1960, pp. 931-932.
5. Pearson, K., "Tables of the Incomplete Gamma Function." Cambridge, The University Press, 1965.
6. Cox, D. R., Miller, H.D., "The Theory of Stochastic Processes." London, Hethuen and Co. Ltd. 1965.

APPENDIX I

TABLE OF INCOMPLETE GAMMA FUNCTION IN THE FORM $P(a, x)$
USED IN CHAPTER III (FROM K. PEARSON [5])

a	x	1	2	4	6	10
1	0.63212	0.86498	0.98168	0.99752	0.99995	
2	0.26424	0.59430	0.90745	0.98257	0.99990	
3	0.08030	0.32362	0.76441	0.93761	0.99959	
4	0.01899	0.14317	0.56653	0.84880	0.99668	
5	0.00366	0.05295	0.37099	0.71374	0.98663	
6	0.00060	0.01686	0.21456	0.55412	0.96183	
7	0.00009	0.00483	0.11026	0.39348	0.91243	
8	0.00002	0.00139	0.05067	0.25579	0.83558	
9	0.00001	0.00027	0.02135	0.15252	0.73226	
10	0.00000	0.00005	0.00813	0.08367	0.60625	
11	0.00000	0.00001	0.00274	0.04139	0.47678	
12	0.00000	0.00000	0.00094	0.02017	0.35680	
13	0.00000	0.00000	0.00028	0.00883	0.24992	
14	0.00000	0.00000	0.00008	0.00356	0.16383	

DATA USED IN EXAMPLE OF CHAPTER III
FORT COLLINS, COLO. STATION NO. 5.3005

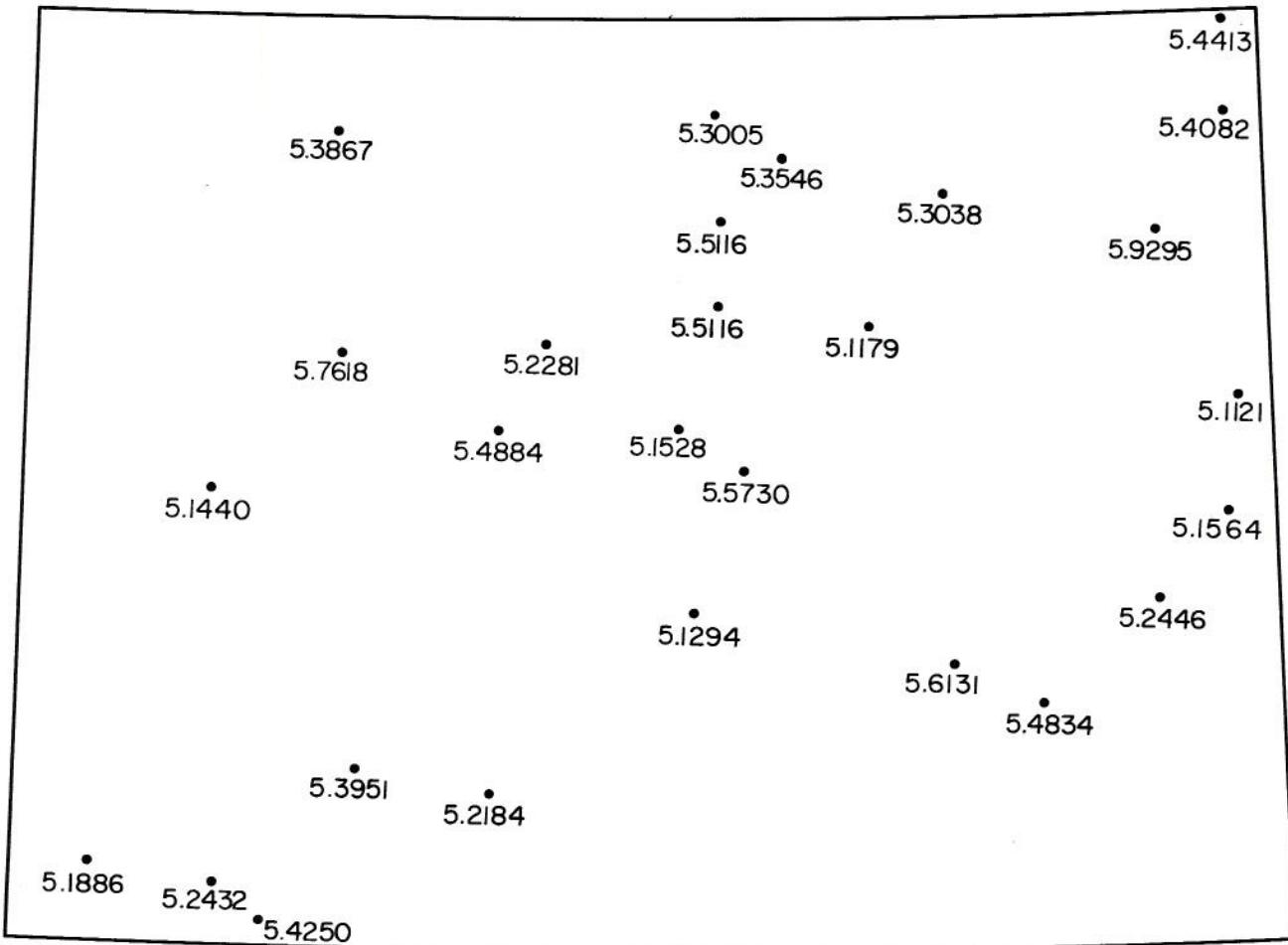
Year	P	S	Year	P	S
1891	17.50	.71	1926	13.57	-.26
1892	13.58	-.26	1927	15.77	.28
1893	5.65	-2.24	1928	13.54	-.27
1894	12.35	-.56	1929	14.08	-.13
1895	18.07	.86	1930	15.17	.13
1896	15.76	.28	1931	9.88	-1.18
1897	15.24	.15	1932	12.80	-.45
1898	11.03	-.89	1933	15.65	.25
1899	16.19	.39	1934	8.87	-1.43
1900	19.21	1.14	1935	15.95	.33
1901	21.17	1.63	1936	11.81	-.70
1902	18.43	.95	1937	12.93	-.42
1903	11.63	-.74	1938	19.72	1.27
1904	13.13	-.37	1939	7.85	-1.69
1905	19.85	1.30	1940	13.94	-.17
1906	19.88	1.31	1941	17.81	.79
1907	11.64	-.74	1942	21.19	1.63
1908	17.22	.64	1943	12.27	-.58
1909	16.24	.40	1944	13.53	-.27
1910	12.92	-.42	1945	15.73	.27
1911	10.89	-.93	1946	14.11	-.12
1912	19.61	1.24	1947	17.95	.83
1913	15.85	.30	1948	10.45	-1.04
1914	14.31	-.07	1949	18.79	1.04
1915	22.79	2.03	1950	12.70	-.48
1916	13.15	-.36	1951	22.52	1.97
1917	13.72	-.22	1952	12.74	-.47
1918	21.79	1.78	1953	11.42	-.80
1919	10.92	-.92	1954	7.98	-1.66
1920	11.65	-.74	1955	12.97	-.41
1921	14.83	.05	1956	12.19	-.60
1922	9.98	-1.16	1957	19.56	1.23
1923	27.57	3.23	1958	17.44	.70
1924	10.64	-.99	1959	14.67	.01
1925	14.46	-.04			

APPENDIX I (continued)

In the preceding Table, P means the total amount of annual precipitation in inches, and S is the total

amount of annual precipitation in standard measure, i.e., $S = \frac{P-\mu}{\sigma}$ where $\mu = 14.62$ and $\sigma = 4.00$.

Precipitation Stations in Colorado

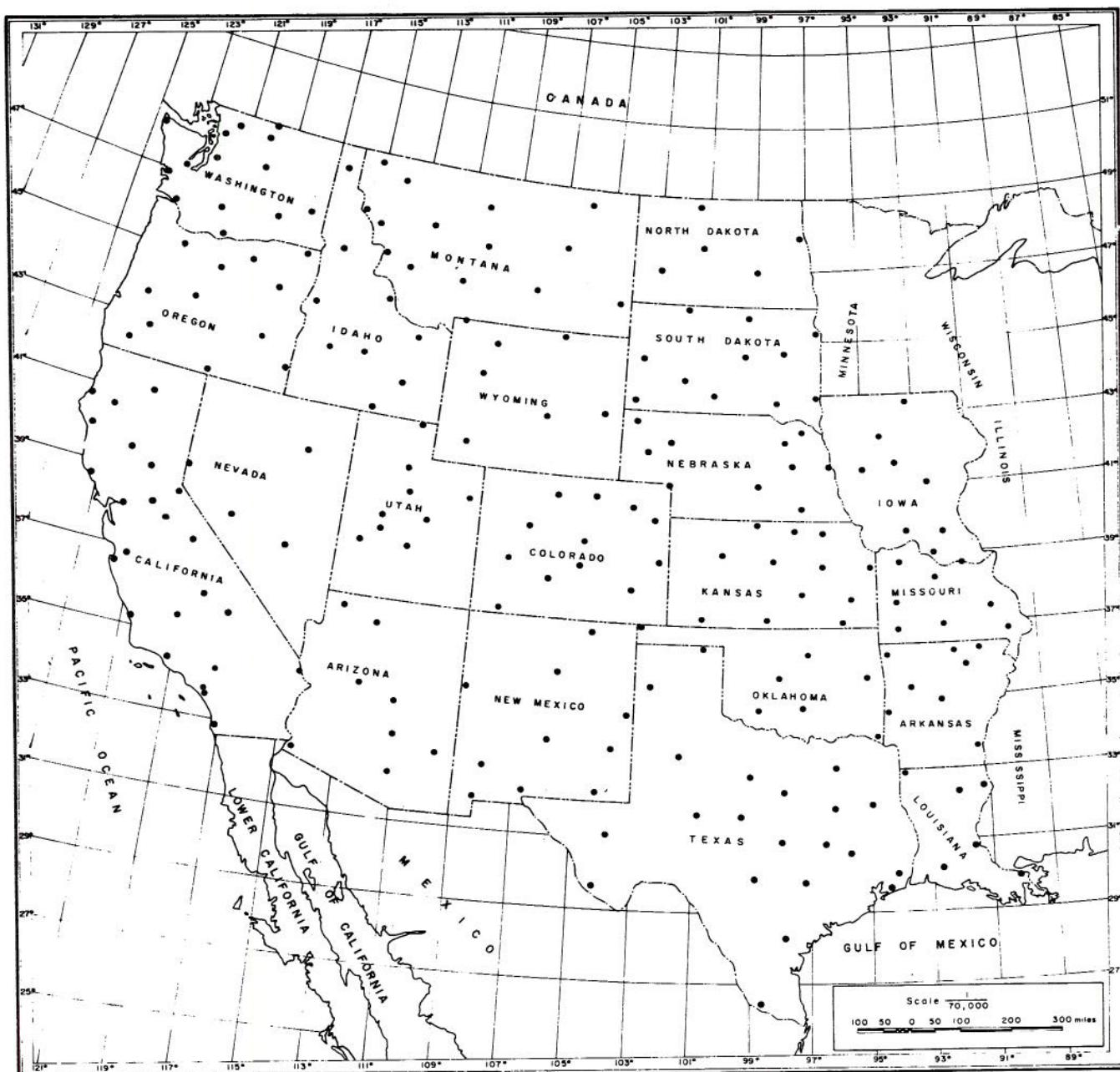


APPENDIX I (continued)

ANNUAL PRECIPITATION
STATIONS OF COLORADO

Station Name	Station No.	N	μ	σ	r	E _(x₁)	E _(x₂)	Var _(x₁)	Var _(x₂)	E(N ₁)	E(N ₂)	Var _(N₁)	Var _(N₂)	E(S ₁)	E(S ₂)	Var _(S₁)	Var _(S₂)	E(I ₁)	E(I ₂)	Var _(I₁)	Var _(I₂)	
Bur-	5.1121	1	67	17.08	4.81	12.1	0.798	0.798	0.363	0.363	2,000	2,000	2,000	2,000	1.5958	1.5958	2,000	2,000	0.798	0.798	0.252	0.252
lington	2					0.735	0.735	0.363	0.363	2,000	2,000	2,000	2,000	1.466	1.466	2,000	2,000	1.880	1.840	0.735	0.160	
Byers	5.1179	1	25	14.04	4.51	9.7	0.819	0.818	0.236	0.413	1,944	1,882	1,599	1,046	1.5917	1.537	1,275	1,262	0.736	0.853	0.149	0.329
	2					0.725	0.725	0.236	0.572	2,000	2,000	2,000	2,000	1.469	1.469	2,000	2,000	1.880	1.840	0.735	0.160	
	3					0.816	0.816	0.235	0.667	2,000	2,000	2,000	2,000	1.469	1.469	2,000	2,000	1.880	1.840	0.735	0.160	
Canon	5.1294	1	36	12.68	3.24	15.3	0.798	0.798	0.363	0.363	2,000	2,000	2,000	2,000	1.5958	1.5958	2,000	2,000	0.798	0.798	0.252	0.252
City	2					0.747	0.747	0.250	0.560	2,149	1,871	2,468	1,628	1.5955	1.368	1,882	1,852	0.740	0.740	0.162	0.416	
3						0.659	0.993	0.363	0.629	3,167	2,429	4,803	1,957	2,050	2,127	2,378	1,311	0.875	0.943	0.318	0.167	
Cedaredge	5.1440	1	33	11.80	2.67	19.5	0.798	0.798	0.363	0.363	2,000	2,000	2,000	2,000	1.5958	1.5958	2,000	2,000	0.798	0.798	0.252	0.252
	2					0.763	0.765	0.550	0.280	1,926	1,783	1,893	1,266	1.5957	1.385	1,892	1,870	0.750	0.750	0.168	0.410	
	3					0.715	0.889	0.163	0.931	2,667	2,429	3,554	5,957	2,0267	2,031	0.659	2,566	0.823	1,060	0.019	1,085	
Cheeseman	5.1528	1	54	15.37	3.15	24.1	0.798	0.798	0.363	0.363	2,000	2,000	2,000	2,000	1.5958	1.5958	2,000	2,000	0.798	0.798	0.252	0.252
	2					0.778	0.778	0.280	0.540	2,117	1,895	2,364	1,677	1.5957	1,408	1,903	1,888	0.750	0.759	0.174	0.400	
	3					0.860	0.413	0.318	1,786	1,933	1,310	3,151	1.5364	1,523	0.936	3,209	1,010	0.223	0.440	0.178		
Cheyenne	5.1564	1	58	16.14	4.81	11.2	0.798	0.798	0.363	0.363	2,000	2,000	2,000	2,000	1.5958	1.5958	2,000	2,000	0.798	0.798	0.252	0.252
Wells	2					0.730	0.730	0.235	0.570	1,973	2,028	1,921	2,081	1.5954	1.5954	2,000	2,000	0.798	0.798	0.252	0.252	
Cortez	5.1886	1	27	15.15	4.29	9.4	0.798	0.798	0.363	0.363	2,000	2,000	2,000	2,000	1.5958	1.5958	2,000	2,000	0.798	0.798	0.252	0.252
	2					0.725	0.725	0.235	0.572	1,911	1,834	2,127	1,806	1.5950	1,540	1,875	1,830	0.727	0.727	0.154	0.428	
	3					0.711	0.866	0.203	0.554	2,023	1,900	2,222	2,050	1.012	1,811	0.502	1,673	0.484	0,881	0.072	0.884	
Del Norte	5.2184	1	32	8.53	2.58	11.0	0.798	0.798	0.363	0.363	2,000	2,000	2,000	2,000	1.5958	1.5958	2,000	2,000	0.798	0.798	0.252	0.252
	2					0.730	0.730	0.238	0.570	1,973	2,028	1,921	2,084	1.5950	1,348	1,878	1,836	0.730	0.730	0.158	0.424	
	3					0.748	0.873	0.202	0.703	1,778	1,778	1,283	1,061	1,3300	1,552	0.867	1,875	0.690	0.836	0.343	0.169	
Dillon	5.2281	1	44	18.23	4.09	19.9	0.798	0.798	0.363	0.363	2,000	2,000	2,000	2,000	1.5958	1.5958	2,000	2,000	0.798	0.798	0.252	0.252
	2					0.765	0.765	0.265	0.550	2,004	1,996	1,989	1,955	1,385	1,893	1,871	1,750	0.750	0.750	0.168	0.410	
	3					0.625	0.612	0.173	0.326	1,357	1,667	1,659	2,088	1,893	1,049	0,671	1,356	0.645	0.613	0.084	0.279	
Durango	5.2432	1	65	19.13	5.49	12.2	0.798	0.798	0.363	0.363	2,000	2,000	2,000	2,000	1.5958	1.5958	2,000	2,000	0.798	0.798	0.252	0.252
	2					0.735	0.735	0.240	0.570	1,778	1,849	2,566	1,569	1,590	1,355	1,880	1,840	0.735	0.735	0.160	0.425	
	3					0.757	0.833	0.219	0.563	2,286	2,067	4,060	2,861	1,7307	1,721	2,321	2,321	0.730	0.730	0.160	0.406	
Eads	5.2446	1	32	13.78	4.19	10.8	0.798	0.798	0.363	0.363	2,000	2,000	2,000	2,000	1.5958	1.5958	2,000	2,000	0.798	0.798	0.252	0.252
	2					0.730	0.730	0.238	0.570	1,973	2,028	1,921	2,084	1.5950	1,348	1,878	1,836	0.730	0.730	0.158	0.424	
	3					0.748	0.873	0.202	0.703	1,778	1,778	1,283	1,061	1,3300	1,552	0.867	1,875	0.690	0.836	0.343	0.169	
Edgewater	5.2557	1	47	15.53	4.37	12.6	0.798	0.798	0.363	0.363	2,000	2,000	2,000	2,000	1.5958	1.5958	2,000	2,000	0.798	0.798	0.252	0.252
	2					0.737	0.737	0.237	0.570	1,568	2,446	1,692	3,735	1,170	1,5950	1,357	1,880	1,736	0.736	0.736	0.160	0.424
	3					0.647	0.666	0.203	0.504	2,000	2,000	2,000	2,000	1.5958	1,5958	2,000	2,000	0.798	0.798	0.252	0.252	
Fort	5.3005	1	69	14.62	4.00	13.4	0.798	0.798	0.363	0.363	2,000	2,000	2,000	2,000	1.5958	1.5958	2,000	2,000	0.798	0.798	0.252	0.252
Collins	2					0.740	0.740	0.245	0,562	1,566	1,857	2,572	1,566	1,5950	1,360	1,880	1,845	0.738	0.738	0.160	0.420	
	3					0.670	0.910	0.243	0,502	1,850	1,524	2,028	0.725	1,239	1,386	0.816	1,027	0,670	0.910	0.123	0.586	
Morgan	5.3038	1	68	13.54	3.50	15.0	0.798	0.798	0.363	0.363	2,000	2,000	2,000	2,000	1.5958	1.5958	2,000	2,000	0.798	0.798	0.252	0.252
	2					0.745	0.745	0.249	0.559	1,529	1,918	2,633	1,532	1,5955	1,368	1,881	1,851	0.740	0.740	0.162	0.416	
	3					0.751	0.783	0.287	0.524	1,834	1,842	1,807	1,767	1,443	1,643	1,006	2,814	0.736	0,685	0.216	0.425	
Greeley	5.3546	1	38	12.16	3.43	12.6	0.798	0.798	0.363	0.363	2,000	2,000	2,000	2,000	1.5958	1.5958	2,000	2,000	0.798	0.798	0.252	0.252
	2					0.737	0.737	0.240	0.568	2,000	2,000	2,000	2,000	1.5958	1.5958	2,000	2,000	0.798	0.798	0.252	0.252	
	3					0.743	0.790	0.267	0.518	2,000	2,000	2,000	2,000	1.5958	1.5958	2,000	2,000	0.798	0.798	0.252	0.252	
Hayden	5.3867	1	28	16.13	3.22	25.0	0.798	0.798	0.363	0.363	2,000	2,000	2,000	2,000	1.5958	1.5958	2,000	2,000	0.798	0.798	0.252	0.252
	2					0.788	0.738	0.243	0.567	2,149	1,871	2,468	1,728	1,5958	1,360	1,880	1,844	0.737	0.737	0.170	0.400	
	3					0.690	0.828	0.246	0.566	2,051	2,051	2,050	2,050	1,5958	1,5958	2,000	2,000	0.798	0.798	0.252	0.252	
Hermit	5.3951	1	38	15.43	4.00	14.9	0.798	0.798	0.363	0.363	2,000	2,000	2,000	2,000	1.5958	1.5958	2,000	2,000	0.798	0.798	0.252	0.252
	2					0.745	0.745	0.249	0.559	1,848	2,185	2,633	1,532	1,5955	1,368	1,881	1,851	0.740	0.740	0.162	0.416	
	3					0.711	0.823	0.236	0.542	2,222	1,800	4,619	1,619	1,560	1,580	1,480	2,279	1,058	0.752	1,035	0.058	0.748
Holyoke	5.4082	1	32	18.12	4.64	15.3	0.798	0.798	0.363	0.363	2,000	2,000	2,000	2,000	1.5958	1.5958	2,000	2,000	0.798	0.798	0.252	0.252
	2					0.747	0.747	0.240	0.557	2,102	1,908	3,215	1,731	1,5955	1,358	1,882	1,741	0.741	0.741	0.163	0.415	
	3					0.801	0.836	0.267	0.435	1,												

APPENDIX II
ANALYSIS OF MONTHLY DATA



Areal distribution of precipitation stations
(After Roesner and Yevjevich)

APPENDIX II (continued)

STATION NAME	STATION NO.	$E(x_1^*)$	$E(x_1^*)^2$	$E(x_2^*)$	$E(x_2^*)^2$	α	β	γ	τ_0	τ_1	$E(N_1)$	Var(N ₁)
CLIFFTON	2-1849	.6670	.5346	1.0371	1.0894	-3.639	.5724	.3886	.0114	2.7452	4.8044	
GRAND CANYON NATIONAL	2-3591	.6658	.5974	.9163	1.4956	+1.01	.5644	.4208	.5792	2.4386	3.5082	
MOUNT TRUMBULL	2-5744	.6671	.5843	.9025	1.4829	+0.98	.5519	.4237	.5763	2.4663	3.6084	
PAYSON RS	2-6320	.6595	.5498	.9571	1.5972	+2.15	.6145	.4069	.5931	2.3725	3.2564	
PINAL RANCH	2-6561	.0232	.4962	.9917	1.7649	+3.716	.5914	.3859	.6141	2.6910	4.5505	
PRESCOTT	2-6796	.6539	.5109	.9675	1.5989	+3.740	.5726	.3951	.6049	2.6738	4.4552	
TUSCON UNIVERSITY OF A	2-8815	.3753	.2126	1.0462	2.0072	+3.37	.5624	.3854	.6552	2.4024	4.0224	
YUMA CITRUS STATION	2-9452	.5953	.2777	1.2787	3.0252	+0.93	.5983	.2830	.5741	2.0114	12.0771	
ARKANSAS CITY	3-1000	.5546	.2777	1.1561	1.5361	+3.98	.5884	.2409	.5991	2.5392	3.9084	
HATSEVILLE LAND O. NV.	3-1060	.6823	.6453	.9238	1.4458	+3.919	.5305	.4249	.5751	2.5915	3.9587	
CUNNAY	3-1596	.6817	.6343	.9129	1.4544	+2.24	.6145	.4280	.5720	2.3616	3.2156	
FAYETTEVILLE EXP. STA	3-2444	.0734	.6088	.8892	1.4840	+3.372	.5762	.4315	.5685	2.2871	2.9437	
MEKA	3-4754	.6659	.6042	.9070	1.4919	+3.335	.5889	.2423	.5760	2.3057	3.0143	
MOUNTAIN HOME I NMW	3-5036	.6735	.6094	.9728	1.5086	+3.374	.5721	.4099	.5901	2.5161	3.8148	
WILLISTON	3-5820	.6679	.6022	.9288	1.5197	+3.375	.5556	.3757	.6125	2.8450	5.2488	
SUBTACO	3-6125	.6626	.6022	.9288	1.5213	+3.465	.5715	.3560	.6040	2.4437	3.5652	
ANTIOCH F. MILLS	3-6227	.6746	.6390	.9998	2.2946	+3.307	.5911	.3344	.6114	2.004	3.3642	
AUDUM	3-6883	.3351	.3964	1.0788	2.1653	+3.097	.6223	.3243	.5677	3.2889	7.1967	
BIG CREEK POWER HOUSE	4-0755	.2739	.3469	.9680	1.5911	+3.067	.6173	.3320	.6640	3.2600	7.3676	
HIG SUR STATE PARK	4-0790	.0092	.3688	1.0375	2.2276	+3.081	.5891	.3714	.6294	2.8729	5.3808	
CHESTER	4-1700	.0227	.4956	.9809	1.7490	+3.208	.6593	.3273	.6727	3.1176	6.6921	
FOOT BRAGG	4-3161	.6659	.4517	1.0195	1.9425	+3.515	.5556	.3757	.6125	2.8450	5.2488	
FOOT POSS	4-3191	.2170	.3694	1.0115	2.1970	+3.294	.5911	.3344	.6114	2.004	3.3642	
HOLISTER	4-3242	.6623	.6022	.9288	1.5213	+3.374	.5911	.3344	.6114	2.004	3.3642	
LYTLE CREEK POWER HNS	4-3251	.2145	.3964	1.1167	2.1653	+3.097	.6223	.3243	.5677	3.2889	7.1967	
MC CLUO	4-5449	.0092	.4489	1.0549	1.7549	+3.270	.6560	.2871	.7129	3.5976	9.3449	
NEEDLES	4-6111	.6633	.2456	1.1484	2.8177	+2.695	.6667	.2879	.7121	3.7101	10.0550	
NEWPORT BEACH HARBOUR	4-6175	.9770	.3007	1.1742	2.0500	+3.004	.7296	.2917	.7083	3.3287	7.7511	
OJAI	4-6399	.0777	.2888	1.1602	2.0659	+3.095	.6556	.3273	.6727	3.1176	6.6921	
SAN DIEGO #4 APT.	4-7740	.6629	.2866	1.0941	2.0500	+3.295	.6643	.2973	.7027	3.3289	8.8474	
SANTA LUIS DRISPO POLY	4-7851	.6651	.2945	1.0778	2.0282	+3.294	.6643	.3016	.6040	2.4437	3.5652	
SUNORA	4-8845	.6643	.5913	.9413	1.5429	+3.397	.5923	.3344	.6114	2.004	3.3642	
TOPANGA PATROL STATION	4-9043	.0043	.4425	1.0003	2.0502	+3.394	.6273	.3677	.6253	2.9437	5.7179	
TRONA	4-9053	.4940	.2952	1.0133	2.0482	+3.295	.6643	.2973	.7027	3.3289	8.8474	
TUSTIN IRVING RANCH	4-9087	.6685	.2910	1.0850	2.0064	+3.294	.6643	.3016	.6040	2.4437	3.5652	
TWIN LAKES	4-9105	.0218	.5062	.9892	1.7174	+3.370	.5890	.3251	.6149	2.3511	3.9084	
WACO	4-9452	.6810	.3195	1.0500	2.5073	+2.986	.6171	.3040	.6057	2.4437	3.5652	
WEAVERTON RS.	4-9464	.0194	.5088	.9797	1.7552	+3.294	.6643	.3016	.6040	2.4437	3.5652	
WILMRS	4-9599	.6625	.3035	1.0503	2.0503	+3.374	.5833	.3349	.6661	3.4249	8.3003	
CANNON CITY	5-1299	.0242	.5107	.9247	1.6878	+3.374	.5879	.3349	.6661	3.4249	8.3003	
CHEESEMAN	5-1524	.6681	.5727	.9707	1.5347	+3.612	.6182	.4272	.5728	2.1685	5.2538	
CHEYENNE MILLS	5-1554	.6602	.4809	.9504	1.7677	+3.632	.5866	.3912	.6088	2.6441	4.3470	
DEL NORTE	5-2184	.6414	.5213	.9543	1.5370	+3.669	.5640	.4132	.5968	2.7111	4.6390	
DURANGO	5-2432	.6653	.5519	.9369	1.5135	+3.610	.5620	.4132	.5968	2.7111	4.6390	
FORT COLLINS	5-3009	.0104	.6907	.9276	1.7424	+3.674	.5922	.3344	.6114	2.004	3.3642	
GOLD MOUNTAIN	5-3218	.6689	.3049	1.1715	2.5564	+2.974	.6116	.3046	.6040	2.4437	3.5652	
JULESBURG	5-3443	.6666	.5911	.9245	1.5921	+3.728	.5893	.3344	.6114	2.004	3.3642	
LAS ANIMAS	5-3484	.0189	.7776	.9822	1.6016	+3.613	.5734	.3344	.6114	2.004	3.3642	
MONROE RD. 2	5-5722	.6673	.6095	.9208	1.4975	+4.091	.5562	.4248	.5762	2.4444	3.5309	
SHOSHONE	5-7511	.0910	.6664	.8811	1.3895	+3.985	.5655	.4247	.5728	2.1685	5.2538	
STEAMBOAT SPRINGS	5-7936	.0724	.5864	.8876	1.3371	+4.047	.5116	.4449	.5581	2.4710	3.6245	
VALLEY	5-9299	.6662	.6062	.9254	1.5771	+3.669	.5640	.4132	.5968	2.7111	4.6390	
ABERDEEN EXPT. STATION	5-9312	.6621	.6088	.9148	1.5370	+3.610	.5620	.4132	.5968	2.7111	4.6390	
ARROWROCK DAM	5-9444	.6612	.5919	.9410	1.5137	+3.669	.5640	.4132	.5968	2.7111	4.6390	
CAMBRIDGE	5-9452	.6685	.2712	1.0500	2.5072	+3.294	.6643	.3016	.6040	2.4437	3.5652	
DUBOIS EXPT. STATION	10-2707	.0471	.5451	.9881	1.5701	+3.294	.6643	.3016	.6040	2.4437	3.5652	
HAILEY RS	10-3042	.6517	.5632	.9556	1.2025	+3.374	.5511	.4047	.5953	2.6691	4.4548	
KOOSKIA	10-5011	.7115	.6826	.9913	1.3542	+4.150	.5216	.4431	.5569	2.4047	3.3970	
OAKLEY	10-6562	.6664	.5823	.9534	1.3066	+3.853	.5562	.4047	.5953	2.5261	3.8387	
SALIN	10-8076	.0055	.6471	.8872	1.3753	+4.000	.5507	.4417	.5583	2.5000	3.7500	
SANDPOINT EXPT. STATION	10-8111	.7200	.6873	.8849	1.3763	+3.905	.5322	.4232	.5768	2.5608	3.9971	
ATLANTIC 1 NE	13-0364	.6815	.6328	.9341	1.4650	+3.294	.6568	.4227	.5773	2.5442	3.9289	
DES MOINES #3 CITY	13-2208	.6717	.6002	.9174	1.3175	+3.930	.5368	.4227	.5773	2.5442	3.9289	

STATION NAME	STATION NO.	$E(x_1^*)$	$E(x_1^*)^2$	$E(x_2^*)$	$E(x_2^*)^2$	α	β	γ	τ_0	τ_1	$E(N_1)$	Var(N ₁)
MASON CITY 3 N	13-5230	.6701	.6240	.9136	1.4646	+2.250	.5743	.4253	.5747	2.3529	3.1834	
OTTUMWA	13-6291	.7114	.6590	.9238	1.4959	+1.413	.5360	.4205	.5650	2.1126	3.4119	
ROCKWELL CITY	13-7161	.6664	.6800	.9213	1.5046	+2.191	.5326	.4205	.5537	2.1276	3.4276	
CONCORDIA #3 CITY	14-1759	.6505	.5596	.9713	1.0292	+3.952	.5900	.4011	.5989	2.5305	3.8771	
COUNCIL GROVE	14-1866	.6517	.5317	.9386	1.3937	+1.301	.5720	.4030	.5970	2.5933	4.1193	
ELLISVILLE	14-2459	.6543	.5727	.9577	1.5756	+1.7594	.5627	.4037	.5921	2.4740	3.7483	
LA CYgne	14-3759	.6502	.5880	.9504	1.5371	+1.374	.5446	.4037	.5813	2.7407	4.7709	
MEDICINE LODGE	14-5173	.6641	.5917	1.0137	.9404	+1.304	.5839	.3846	.6134	2.7407	4.7709	
PHILLIPSBURG	14-6374	.6543	.5215	.9591	1.5890	+1.629	.5864	.3988	.6173	2.5797	4.0377	
PLAINS	14-6647	.6633	.5776	.9314	1.5856	+1.037	.5592	.4167	.5833	2.7099	3.6582	
QUINN	14-6657	.6633	.5212	.9591	1.5871	+1.7070	.5844	.3864	.6134	2.7407	4.7709	
SEJAN	14-6857	.6648	.5655	.9564	1.3659	+1.374	.5878	.3874	.6080	2.8045	3.7684	
SEGUIN	14-7213	.6674	.5868	.9942	1.5800	+1.0798	.5871	.3874	.6134	2.7407	4.7709	
CALHOUN EXP. STATION	15-1411	.6551	.5945	.9125	1.3921	+1.374	.5474	.3846	.6134	2.7407	4.7709	
JENNINGS	15-6700	.0239	.5030	.9225	1.6501	+2.227	.5705	.4034	.5966	2.5940	4.1263	
HEMANN	15-6811	.6549	.5942	.9244	1.2247	+1.374	.5474	.3846	.6134	2.7407	4.7709	
SHIELDINA	15-7393	.6718	.6040	.9547	1.5748	+1.0795	.5844	.3864	.6134	2.7407	4.7709	
WAHWEEDJAH	15-8823	.0234	.5777	.9633	1.3744	+1.374	.5376	.3874	.6134	2.7407	4.7709	
WILLOW SPRINGS	15-											

APPENDIX II (continued)

STATION NAME	STATION NO.	$E(x_1^*)$	$E(x_1^*)^2$	$E(x_2^*)$	$E(x_2^*)^2$	α	β	π_0	π_1	$E(N_1)$	Var(N_1)
PECOS RS	29.6676	.0591	.5660	.9313	1.3461	*154	.5847	.4193	.5947	2.4974	3.3882
STATE UNIVERSITY	29.8539	.0914	.4231	.7885	1.7252	.5610	.5970	.3721	.6210	2.1257	4.9046
ZUMA FALLS	29.8540	.0933	.4231	.7885	1.5679	.5615	.5971	.3721	.6210	2.1257	4.9046
DICKENSON CAPT. STATION	32.1840	.5389	.5344	.9045	1.3740	*170	.5663	.4010	.5990	2.6363	3.4223
GRAND FORKS U.	32.3621	.0476	.5584	.9066	1.5824	*203	.5872	.4172	.5828	2.3793	3.2818
JAMESIOMM ST., MOSP.	32.4418	.5375	.5537	.9730	1.5651	*1793	.5776	.3964	.6036	2.6364	4.3140
MAX	32.5638	.0915	.6392	.8841	1.1833	*455	.5669	.4001	.5599	2.2444	2.7931
MUHALL	32.6025	.8701	.6041	.9575	1.5295	*144	.5934	.4112	.5888	2.133	3.1016
NEEDELL	32.6115	.0831	.6261	.9773	1.2525	*216	.5872	.4152	.5825	2.3714	3.2716
ISABEL	34.4451	.0615	.4241	.7931	1.2525	*216	.5575	.4300	.5700	2.3774	3.2752
KENTON	34.4766	.0603	.4336	1.0063	1.1746	*314	.5932	.3933	.6047	2.9495	6.4524
PAULS VALLEY	34.6926	.0629	.5668	.9961	1.3556	*051	.5561	.4232	.5768	2.5488	3.2490
PENNY	34.7012	.0649	.5861	.9287	1.3721	*272	.6184	.4079	.5966	2.347	3.1380
WEEDER FALLS	34.9445	.0601	.5860	.9242	1.3721	*272	.5934	.4113	.5827	2.5467	3.1562
WICHITA MT. ALH	34.9610	.0709	.5224	.9393	1.3721	*272	.5865	.4137	.5827	2.347	3.1380
ANADALE 1 N	35.0497	.0714	.5556	.8962	1.3863	*130	.5514	.4292	.5718	2.4216	3.2218
BEND	35.0694	.0295	.102	.9915	1.1700	*3924	.6197	.3477	.6143	2.5466	3.1468
CUTTAW GROVE 1 S	35.1997	.0625	.5124	.9203	1.1872	*374	.5520	.4195	.5814	2.5104	3.1158
DANNER	35.2135	.0735	.6070	.9429	1.1702	*3667	.5133	.4167	.5833	2.7273	4.1107
ESTACADA & SE	35.2691	.0615	.5700	.9306	1.3971	*050	.5488	.4236	.5745	2.4490	3.4268
WHITE PASS	35.3445	.0644	.5371	.9373	1.3721	*272	.5865	.4139	.5827	2.347	3.1380
MONPENIS	35.4471	.0724	.5754	.9752	1.2525	*216	.5814	.4301	.5768	2.3486	3.1380
LAKEVIEW	35.4470	.0249	.5250	.9671	1.6774	*3811	.5859	.3941	.6059	2.021	4.2616
MINAM 7 NE	35.5610	.0906	.6585	.9148	1.4019	*3833	.5038	.321	.5679	2.6090	4.1980
PROSPECT 2 SW	35.6907	.0323	.5751	.9206	1.3761	*4920	.5500	.4072	.5928	2.6101	4.2382
ROCK CREEK	35.7250	.1700	.5908	.9036	1.3395	*109	.5217	.4211	.5949	2.4350	3.4899
WALNUT SPRINGS RESERVAIR	35.9044	.0675	.5673	.9238	1.3175	*212	.5377	.4124	.5800	2.3126	3.5226
ARMOUR	35.9142	.0712	.5712	.9242	1.3721	*3792	.5367	.4148	.5866	2.4309	3.4899
CUTTAW 4000	35.9172	.0264	.5077	.9418	1.2523	*2816	.5859	.3879	.6021	2.7098	4.8849
EUREKA	39.2797	.0274	.5324	.9385	1.5616	*3770	.5622	.401	.5986	2.6549	4.3832
HIGHMIRE 1 W	39.3832	.0510	.5620	.9672	1.0078	*053	.6022	.4024	.5971	2.4615	3.5976
HOT SPRINGS	39.4007	.0187	.5201	.9372	1.5687	*4447	.5228	.3941	.6033	2.4101	5.5168
LAUREL 7 NE	39.4461	.0712	.5711	.9712	1.5629	*3824	.5859	.4064	.5914	2.5507	3.611
LEHMONT	39.4464	.0194	.5171	.9261	1.3736	*3567	.5940	.3709	.6234	2.7801	4.9556
MILBANK	39.5536	.0646	.5769	.9183	1.2041	*3928	.5592	.4143	.5857	2.5499	3.4358
SIOUX FALL #8 AP	39.7667	.0702	.6255	.9448	1.1676	*0494	.5629	.4114	.5816	2.4697	3.6297
VALE	39.8552	.0144	.5051	.9198	1.0943	*3957	.5904	.4113	.5987	2.5270	3.8588
VOO	39.9442	.0255	.5223	.9931	1.5885	*3746	.5997	.4049	.5814	2.4309	4.3094
ALLANT	41.0120	.0702	.5202	.9010	1.3721	*3746	.5860	.4074	.5924	2.4013	3.2211
BALTIMORE EXP. STATION	41.0404	.0008	.4445	1.0062	1.6656	*3746	.5576	.3750	.6250	2.3492	5.4566
BEAUMONT	41.0411	.0733	.5504	.9295	1.5265	*021	.5879	.4001	.5939	2.4472	3.6989
BUONARDOO	41.1138	.0720	.5209	1.0000	1.7200	*3700	.5824	.3873	.6127	2.7067	4.6019
CORSICANA	41.2019	.0508	.5908	.9113	1.3549	*3852	.5442	.4146	.5856	2.5901	4.1435
FLATIRON	41.3182	.0423	.5350	.9211	1.3721	*3746	.5616	.4049	.5930	2.4774	3.2745
GALVESTON #3 CITY	41.3404	.0701	.5620	.9626	1.1706	*2747	.5745	.4292	.5892	2.6661	4.3332
GEORGES WEST	41.3504	.0049	.4549	1.0368	1.0833	*1754	.6414	.3892	.6308	2.6661	4.3332
GREENVILLE 2 SW	41.3733	.0319	.5360	.9532	1.0556	*3850	.5801	.3989	.6011	2.5976	4.1948
HENDERSON	41.4081	.0703	.6200	.9021	1.6663	*2747	.5734	.4270	.5730	2.3349	3.1351
KERRVILLE	41.4780	.0260	.5016	.9192	1.6944	*3750	.5506	.4051	.5949	2.6667	4.4444
LAMARAS	41.5010	.0438	.5404	.9494	1.3370	*3800	.5730	.4033	.5927	2.4737	4.1744
MISSISS	41.5072	.0137	.5177	.9240	1.4824	*3250	.5849	.4071	.5927	2.5229	3.6742
PERRYTON	41.5173	.0274	.5214	.9211	1.2445	*3859	.5745	.4040	.5920	2.5229	3.6544
POST	41.7204	.0260	.5304	1.0048	1.6865	*0337	.6429	.3738	.6262	2.6063	4.1865
PRESIUDIO	41.7252	.0718	.4037	1.1018	2.4051	*3291	.6371	.3406	.6594	3.3465	6.1938
RIVERSIDE	41.7651	.0113	.5968	.9314	1.2297	*0509	.5669	.4159	.5841	2.4571	3.6275
STERLING CITY	41.8630	.0351	.5975	1.0097	1.5997	*3756	.5111	.3801	.6111	2.4411	3.2899
VALLEY JUNCTION	41.9230	.0102	.5350	.9211	1.3721	*3746	.5182	.4124	.5876	2.7530	4.7006
VEGA	41.9230	.0102	.5774	.9367	1.1704	*3670	.5603	.3958	.6042	2.7250	4.7006
WEATHERFORD	41.9332	.0277	.5225	.9222	1.0649	*3674	.5501	.4049	.5951	2.7216	4.6955
DESSEM	42.2101	.0535	.5387	.9305	1.6171	*1631	.5927	.4126	.5874	2.4042	3.4685
SPANISH FORK POWER HOU	42.8119	.7038	.6732	.8957	1.3704	*4286	.5472	.4392	.5608	2.3333	3.1111
TOOELE	42.8771	.7066	.6702	.8802	1.3727	*4531	.5643	.4453	.5547	2.2073	2.6647
BROOKLYN	45.0917	.7238	.7046	.8846	1.2869	*4242	.5215	.4486	.5514	2.3571	3.1990
CEDAR LAKE	45.1223	.7214	.7351	.8331	1.2657	*4638	.5370	.4634	.5366	2.1561	2.4926
CHELAN	45.1350	.6446	.5335	.9679	1.6642	*3783	.5680	.3998	.6002	2.6436	4.3451
COLFAX 1 NW	45.1586	.6209	.5511	.9370	1.6400	*3855	.5818	.3986	.6014	2.5938	4.1338
GOLDENDALE	45.3222	.6310	.5333	.9665	1.6641	*3912	.5992	.3950	.6050	2.5563	3.9785
HATTON 8 E	45.3546	.6279	.5304	.9291	1.6596	*3840	.5699	.4026	.5974	2.6039	4.1764
LONGVIEW	45.4769	.7467	.7454	.8582	1.2349	*4589	.5297	.4662	.5358	2.1792	2.5659
NEWHAILEM	45.5840	.7471	.7634	.8586	1.2123	*4502	.5198	.4641	.5359	2.2212	2.7124
RIMROCK TETON DAM	45.7038	.6392	.5382	.9256	1.6207	*3812	.5520	.4085	.5915	2.6232	4.2579
SEDRO WOLLEY 1 E	45.7507	.7057	.7074	.8790	1.3224	*4437	.5526	.4453	.5547	2.2540	2.8264
SHELTON	45.7584	.7031	.6724	.8789	1.3345	*4600	.5750	.4444	.5556	2.1739	2.5520
SUNNYSIDE	45.8207	.6276	.5108	1.0021	1.7418	*3655	.5850	.3845	.6155	2.7360	4.7495
TATOOSH ISLAND WB	45.8332	.7063	.7033	.8894	1.3493	*4350	.5488	.4421	.5579	2.2991	2.9868
WINTHROP 1 WSW	45.9376	.6599	.5520	.9320	1.5709	*3832	.5412	.4145	.5855	2.6095	4.2001
BUFFALO BILL DAM	46.1175	.6341	.5135	.9335	1.6667	*3965	.5855	.4038	.5962	2.5221	3.8387
DUBOIS	46.2715	.6449	.5385	.9291	1.6140	*3706	.5359	.4088	.5912	2.6984	4.5830
GREEN RIVER	46.4065	.6373	.5488	.9644	1.6178	*3841	.5812	.3979	.6021	2.6036	4.1751
LUSK	46.5830	.6558	.5874	.8876	1.5179	*3958	.5316	.4268	.5732	2.5266	3.8572
PATHFINDER DAM	46.7105	.6567	.5662	.9038	1.5573	*4245	.5842	.4208	.5792	2.3559	3.1945
SHERIDAN FIELD STATION	46.8160	.6465	.5651	.9345	1.5718	*4033	.5849	.4081	.5919	2.4797	3.6691
YELLOWSTONE PARK	46.9905	.6173	.5314	.9109	1.6571	*3612	.5330	.4039	.5961	2.7688	4.8975

APPENDIX II (continued)

STATION NAME	STATION NO.	EN ₂	Var N ₂	ES ₁	Var S ₁	ES ₂	Var S ₂	EI ₁	Var I ₁	EI ₂	Var I ₂
CLIFTON	2-1849	1.7464	-1.3935	1.4096	-2.3561	1.8111	-2.6770	.6570	.054	-1.0371	-.455
GRAND CANYON NATIONAL MOUNT TRUMBULL	2-3591	1.7771	-1.3978	1.4224	-1.9341	1.6221	-2.3930	.6558	.055	.0163	.455
PAYSON HS	2-5744	1.6788	-1.4211	1.5539	-1.4931	1.6551	-2.4103	.6548	.0545	.0225	.4901
ELKO RANCH	2-6320	1.6275	-1.0211	1.5577	-1.6540	1.5577	-2.4440	.6545	.0541	.0571	.5287
PREScott	2-6796	1.6910	-1.3885	1.6769	-2.0574	1.6769	-2.6455	.6232	.0532	.0917	.5918
TUSCON UNIVERSITY OF A YUMA CITRUS STATION	2-8815	1.7465	-1.3937	1.6949	-2.0598	1.6998	-2.5510	.6339	.0541	.0675	.5691
ARKANSAS CITY BATESVILLE LAND D+ NO.	2-9652	1.6529	-1.0923	1.7292	-2.2976	1.7292	-2.6932	.5753	.0497	1.0462	.7041
CUNWAY FAYETTEVILLE EXP. STA	3-0234	1.4882	-1.1888	1.5945	-1.9483	1.5977	-2.3224	.6280	.0570	.9401	.5673
WENDELL HENRY MOUNTAIN HOME 1 NW POCATONAS SUBIACO	3-0460	1.6888	-1.6578	1.7049	-2.2445	1.7409	-2.5320	.6823	.1085	.0336	.4219
ANTIOCH F. MILLS AUBURN BIG CREEK POWER HOUSE	3-1593	1.7663	-1.3948	1.6978	-1.4948	1.6891	-2.2510	.6171	.0541	.0129	.4446
BIG SUR STATE PARK CHESTER FORGEAGG FURT ROSS	3-2444	1.6566	-1.3947	1.5844	-1.6891	1.6432	-2.1248	.6734	.0566	.8892	.5197
LITTLE CREEK POWER HOUS MC CLOUD NEEDLES NEWPORT BEACH MARSH OJAI SAN DIEGO WB APT.	3-4236	1.7840	-1.3747	1.6964	-2.1224	1.7005	-2.2201	.6735	.0598	.9728	.4193
MULLISTER TOWNE LITTLE CREEK POWER HOUSE	3-5820	1.7056	-1.2334	1.5798	-1.7998	1.5831	-2.1525	.6698	.0570	.92d2	.4950
SUBIACO ANTIOCH F. MILLS	3-6928	1.7527	-1.3194	1.6168	-1.9283	1.6168	-2.2275	.6733	.1049	.9224	.4700
AUBURN BIG CREEK POWER HOUSE	4-0227	1.6070	-1.954	1.6067	-2.0770	1.6067	-3.0502	.4976	.0560	.9998	1-0092
4-0383	1.8200	-1.0844	1.7444	-2.4044	1.7444	-2.7870	.5351	.053	1.018	.7827	
4-0755	1.6955	-1.3925	1.6251	-2.0810	1.6251	-2.8272	.5350	.0505	.9860	.4331	
4-0790	1.6710	-1.7836	1.5780	-2.024	1.5735	-2.5889	.5662	.0504	1.0375	.9280	
4-1000	1.6200	-1.0000	1.7717	-2.3422	1.7558	-2.7910	.6227	.0511	.9809	.5737	
4-1110	1.6273	-1.0209	1.6631	-2.2163	1.6591	-2.5234	.5649	.0736	1.0195	.7018	
4-1311	1.5926	-1.958	1.6082	-2.0802	1.6110	-2.8383	.5177	.0549	1.0115	.9234	
4-4022	1.5209	-7.723	1.5847	-2.1337	1.5821	-2.9086	.4723	.0566	1.0402	.0855	
4-5215	1.4057	5.756	1.5775	-1.9418	1.5743	-2.7932	.4925	.0380	1.1167	1-0455	
4-5449	1.5884	-1.1824	1.6122	-1.9432	1.6122	-2.9117	.5662	.0571	.9553	.4553	
4-6118	1.5516	-1.626	1.6171	-2.1492	1.6171	-2.8122	.5653	.0513	.9884	.4194	
4-6715	1.5000	-7.700	1.7698	-2.5593	1.7613	-2.7950	.4770	.0554	1.1742	.9553	
4-7370	1.5706	-5.980	1.5902	-1.9707	1.5902	-2.9249	.4777	.0513	1.1602	.1223	
4-7740	1.4613	-6.740	1.5987	-2.0505	1.5979	-2.9388	.6269	.0365	1.0941	.1998	
4-7851	1.5640	-6.792	1.5778	-2.0509	1.5778	-2.9358	.4567	.0343	1.0778	.2047	
4-8045	1.6600	-1.0756	1.5625	-1.8056	1.5625	-2.9355	.6152	.0775	.9413	.6119	
4-8333	1.5692	-8.7933	1.5798	-1.9798	1.5798	-2.9151	.5628	.0561	1.0155	.8315	
4-8867	1.4933	-1.687	1.6120	-2.2413	1.6120	-2.9151	.5628	.0561	1.0155	.4559	
4-9035	1.4829	-7.7161	1.6190	-2.0895	1.6190	-2.9151	.4987	.0560	1.033	.0443	
4-9105	1.6698	-1.1885	1.6080	-2.0371	1.6617	-2.9283	.6218	.0704	.9892	.5657	
4-9452	1.6204	-1.0505	1.7907	-2.0731	1.7907	-2.9152	.6810	.0425	1.1050	1-0022	
4-9490	1.7606	-1.3392	1.7289	-2.2454	1.7249	-2.8823	.6194	.0572	.9797	.5300	
4-9699	1.7316	-1.2988	1.8371	-2.1619	1.8371	-2.9150	.6194	.0533	1.1633	.7918	
5-1294	1.7609	-1.3398	1.6283	-1.9260	1.6283	-2.6160	.6242	.0511	.9247	.6162	
5-1528	1.6989	-1.2734	1.6198	-1.9237	1.6198	-2.6160	.6242	.0511	1.0707	.6339	
5-1564	1.5989	-1.1873	1.6108	-1.9236	1.6108	-2.6160	.6242	.0511	.9504	.6571	
5-1813	1.5228	-1.5926	1.5926	-1.9442	1.5926	-2.6160	.6242	.0511	.9543	.5301	
5-2432	1.7747	-1.3479	1.6622	-1.9443	1.6622	-2.6162	.6242	.0511	.9369	.5447	
5-3005	1.7471	-1.3052	1.6333	-1.9405	1.6333	-2.6157	.6101	.0596	.9376	.6448	
5-3038	1.6542	-1.5838	1.6501	-2.0255	1.6501	-2.6161	.6237	.0741	.8894	.6186	
5-4413	1.6760	-1.6333	1.7344	-2.1361	1.7344	-2.7934	.6486	.0768	.9745	.5317	
5-4534	1.7080	-1.2948	1.7080	-1.9442	1.7080	-2.6161	.6237	.0511	.9282	.4774	
5-5722	1.9779	-1.9443	1.7428	-2.2942	1.7428	-2.7124	.6101	.0511	.8811	.4294	
5-5723	1.9568	-1.8724	1.7430	-2.3000	1.7430	-2.7124	.6101	.0511	.8876	.3858	
5-6925	1.6299	5.1816	1.6594	-2.0348	1.6593	-2.6230	.6642	.0720	.9254	.5263	
5-7205	1.6866	-1.1579	1.5378	-1.6235	1.5378	-2.3573	.6322	.0666	.9118	.6296	
10-0010	1.8897	1.6047	1.7662	-2.1745	1.7538	-2.7481	.6512	.0375	.9510	.4965	
10-0446	1.7747	-1.3479	1.6622	-2.0142	1.6622	-2.7207	.6614	.0541	.9543	.5301	
10-1408	1.7180	-1.2288	1.6827	-2.0340	1.6827	-2.7207	.6614	.0541	.9369	.5447	
10-2707	1.6730	-1.0505	1.6922	-2.0996	1.6922	-2.7207	.6785	.0541	.9891	.5333	
10-3163	1.6163	-1.7773	1.739	-2.2082	1.739	-2.6359	.6517	.0697	.9556	.5188	
10-5011	1.9172	-1.7968	1.7146	-2.1449	1.7146	-2.7059	.6733	.0533	.9392	.3972	
10-6542	1.7978	-1.6453	1.7141	-2.1855	1.7141	-2.6861	.6000	.0874	.9434	.4835	
10-8076	1.7416	-1.2710	1.5449	-1.6236	1.5449	-2.0562	.7055	.0984	.8872	.4469	
10-8137	1.9776	-1.9331	1.7499	-2.2807	1.7499	-2.6851	.7000	.1083	.8849	.4131	
10-8384	1.8289	-1.6515	1.7084	-2.2299	1.7084	-2.7582	.6585	.0965	.9347	.4247	
13-0384	1.8620	1.6073	1.7090	-2.1519	1.7090	-2.6118	.6717	.0901	.9174	.4873	

STATION NAME	STATION NO.	EN ₂	Var N ₂	ES ₁	Var S ₁	ES ₂	Var S ₂	EI ₁	Var I ₁	EI ₂	Var I ₂
MASON CITY 3 N DTUWA RUCKEWS CITY	13-5230	1.7412	1-2705	1.6908	1-8793	1.5908	2-1792	.6761	.1055	.9136	-.7335
13-6391	1.6585	1-2575	1-7170	2-0456	1-7170	2-1946	.7114	.055	.9238	-.4111	
13-7120	1.6585	1-2575	1-7170	2-0456	1-7170	2-1946	.7114	.055	.9238	-.4111	
14-1759	1.6988	1-1776	1.6642	-1.9443	1.6642	-2.6445	.6503	.0945	.9225	.5169	
14-1866	1.7403	-1.0211	1.6301	-1.6202	1.6301	-2.6755	.6317	.0935	.9347	.5544	
14-2455	1.6897	-1.6811	1.6047	-2.4005	1.6047	-2.7880	.6593	.0741	.9337	.4426	
14-3759	1.8929	1-5175	1.7295	-2.2146	1.7295	-2-5962	.6562	.0814	.9436	.4426	
14-4022	1.6153	-1.6227	1.6153	-2.1721	1.6153	-2-5962	.6563	.0814	.9436	.4426	
14-4613	1.6171	-1.1415	1.6171	-1.9434	1.6171	-2-5962	.6564	.0814	.9436	.4426	
14-5374	1.6374	-1.3027	1.6171	-1.9434	1.6171	-2-5962	.6565	.0814	.9436	.4426	
14-6374	1.7592	-1.3008	1.6171	-1.9434	1.6171	-2-5962	.6565	.0814	.9436	.4426	
14-6427	1.8195	-1.4912	1.7766	-2.3444	1.7766	-2-5962	.6565	.0814	.9436	.4426	
14-6637	1.7024	-1.1575	1.6110	-1.9434	1.6110	-2-5962	.6566	.0814	.9436	.4426	
14-7305	1.8549	-1.0337	1.7024	-2.2165	1.7024	-2-5962	.6567	.0814	.9436	.4426	
14-7311	1.8235	-1.5175	1.6171	-1.9434	1.6171	-2-5962	.6568	.0814	.9436	.4426	
14-7326	1.6551	-1.5225	1.6706	-2.2119	1.6706	-2-5962	.6569	.0814	.9436	.4426	
14-7745	1.6171	-1.3107	1.6171	-1.9434	1.6171	-2-5962	.6570	.0814	.9436	.4426	
14-7747	1.6171	-1.3107	1.6171	-1.9434	1.6171	-2-5962	.6571	.0814	.9436	.4426	
14-7750	1.6171	-1.3107	1.6171	-1.9434	1.6171	-2-5962	.6572	.0814	.9436	.4426	
14-7752	1.6171	-1.3107	1.6171	-1.9434	1.6171	-2-5962	.6573	.0814	.9436	.4426	
14-7754	1.6171	-1.3107	1.6171	-1.9434	1.6171	-2-5962	.6574	.0814	.9436	.4426	
14-7757	1.6171	-1.3107	1.6171	-1.9434	1.6171	-2-5962	.6575	.0814	.9436	.4426	
14-7759	1.6171	-1.3107	1.6171	-1.9434	1.6171	-2-5962	.6576	.0814	.9436	.4426	
14-7761	1.6171	-1.3107	1.6171	-1.9434	1.6171	-2-5962	.6577	.0814	.9436	.4426	
14-7763	1.6171	-1.3107									

APPENDIX II (continued)

STATION NAME	STATION NO.	EN ₂	Var N ₂	ES ₁	Var S ₁	ES ₂	Var S ₂	EI ₁	Var I ₁	EI ₂	Var I ₂
PEOCO RS	29.4676	1.71102	1.2148	1.5867	1.7886	1.5929	2.2627	.6591	.0921	.0313	.51x3
STATE UNIVERSITY	29.8539	1.67160	1.1453	1.6307	1.1455	1.6408	2.7134	.5914	.0422	.0766	.7412
ZUNI FAA AP.	29.9897	1.64818	1.6045	1.7629	2.2336	2.6901	.6538	.0724	.0433	.0433	.4491
DICKENSON EXP. STATIO	32.2188	1.7560	1.3226	1.6856	2.0774	1.6856	2.5539	.6389	.0747	.0555	.5555
GRAND FORKS U.	32.3621	1.7039	1.1771	1.5403	1.7712	1.5439	2.2807	.6476	.0874	.0966	.5765
JAMESTOWN ST. HOSP.	32.4418	1.7312	1.2958	1.6860	2.1415	1.6844	2.4059	.6375	.0873	.0730	.5234
KAN	32.5038	1.7540	1.3149	1.5520	1.9770	1.5552	2.3415	.6415	.1046	.0844	.5435
MUHALL	32.5620	1.7537	1.3057	1.5916	1.9816	1.6001	2.3014	.6515	.1046	.0876	.4466
GEARY	32.6027	1.7567	1.3111	1.5831	1.9824	1.5881	2.0272	.6618	.0909	.0928	.4224
KENTON	34.4457	1.7938	1.0440	1.6202	1.9101	1.6276	2.2546	.6815	.1029	.0973	.4442
PAULS VALLEY	34.4766	1.6859	1.1764	1.6099	2.1871	1.6999	2.7298	.5803	.0540	.0083	.7020
PARK	34.6926	1.6113	1.0496	1.6386	1.9978	1.6412	2.5281	.6629	.0783	.0661	.5343
WEHNER FALLS	34.8112	1.6222	1.0500	1.5902	1.0115	1.5930	2.4054	.6701	.0871	.0871	.5525
WICHITA MT. WLR	34.9155	1.6537	1.2467	1.6535	1.9425	1.6525	2.3341	.6601	.0740	.1042	.5195
ANTELOPE 1 N	35.0197	1.6137	1.0459	1.6383	1.9222	1.6383	2.2528	.6709	.0881	.0593	.4747
BEND	35.0934	1.6138	1.0905	1.6043	1.9944	1.6001	2.4262	.6712	.0864	.0902	.5014
COTTAGE GROVE 1 C	35.1897	1.6115	1.0700	1.6071	2.1113	1.6671	2.3993	.6625	.1054	.0203	.4465
DAMON	35.2135	1.6289	1.1590	1.6770	2.4758	1.6819	2.8273	.6915	.0891	.0884	.4888
ESTACADA 2 SE	35.2393	1.6289	1.1590	1.6770	2.4758	1.6819	2.8273	.6915	.0891	.0884	.4888
GRANTS PASS	35.3445	1.6567	1.1111	1.6301	2.0166	1.6301	2.4417	.6747	.0761	.0334	.5334
HOPPER	35.3827	1.7197	1.2378	1.6770	2.9897	1.6770	2.1304	.6751	.1003	.0752	.4175
LAKEVIE	35.4670	1.7068	1.2063	1.6584	2.0257	1.6508	2.3948	.6290	.0763	.0671	.5513
MINNEAPOLIS NE	35.5610	1.9851	1.0954	1.6228	2.0396	1.6159	2.7582	.6986	.1014	.0148	.3934
PROSPECT 2 SW	35.6907	1.7085	1.0494	1.6530	2.0750	1.6530	2.0225	.6705	.1042	.0206	.5325
ROCK CREEK	35.7040	1.6951	1.0576	1.6591	2.1740	1.6591	2.4414	.6709	.1042	.0206	.4888
WARM SPRINGS RESERVOIR	35.9046	1.6833	1.0705	1.6807	2.0397	1.6807	2.4837	.6675	.0741	.0238	.4464
AKHOUR	39.0296	1.6531	1.0800	1.7592	2.2869	1.7592	2.6525	.6671	.0842	.0442	.4717
COTTONWOOD	39.1972	1.6284	1.5145	1.7278	2.4242	1.7119	2.6179	.6244	.0878	.0418	.5874
EURYKA	39.2797	1.7768	1.3857	1.6063	2.1116	1.6594	2.5723	.6274	.0859	.0366	.5823
HIGHGROVE 1 N	39.3857	1.6697	1.0900	1.6255	1.0502	1.6552	2.1552	.6611	.0852	.0572	.5161
HOT SPRINGS	39.4077	1.6536	1.0575	1.6307	2.0156	1.6307	2.4417	.6747	.0761	.0311	.5334
LAUVELLE 7 NE	39.4461	1.7002	1.1709	1.6530	1.9428	1.6530	2.4084	.6168	.0878	.0513	.5130
LEMMON	39.4864	1.6856	1.1576	1.6228	2.1268	1.6508	2.3948	.6290	.0763	.0671	.5513
MILBANK	39.5538	1.6610	1.0427	1.6538	2.0597	1.6538	2.5148	.6694	.0936	.0183	.5293
SIOUX FALLS NW AP	39.7667	1.7766	1.3798	1.6749	2.0783	1.6749	2.2887	.6782	.1119	.0448	.4402
VALLEY	39.8552	1.6939	1.1515	1.5537	1.7600	1.5551	2.4425	.6148	.0772	.0398	.6225
WUOD	39.9155	1.7005	1.2480	1.6749	2.0783	1.6749	2.3234	.6740	.0901	.0307	.4737
ALDANY	41.0120	1.7421	1.2728	1.6264	1.3055	1.6376	2.4801	.6454	.0843	.0420	.5744
BALMORHEA EXP. STATION	41.0498	1.7935	1.1621	1.7930	2.4375	1.8043	2.9862	.6008	.1044	.0102	.6762
BEAUMONT	41.0611	1.7010	1.1295	1.5777	1.9851	1.6051	2.4205	.6343	.0707	.0295	.5749
BIGFOOT	41.1138	1.7081	1.2095	1.7042	2.1664	1.7062	2.4394	.6320	.0709	.1000	.5443
COHICANA	41.2019	1.6837	1.1547	1.6250	2.1255	1.6250	2.4052	.6592	.0934	.0310	.4930
FLATONIA	41.2413	1.6968	1.1942	1.6270	1.9694	1.6270	2.1213	.6592	.0841	.0111	.5117
GALVESTON WB CITY	41.3430	1.7047	1.2094	1.6209	2.0069	1.6767	2.5584	.6309	.0924	.0620	.5864
GEORGE WEST	41.3504	1.5591	1.0116	1.6115	1.8725	1.6161	2.1959	.6049	.0553	.0366	.6282
GREENVILLE 2 SW	41.3734	1.7239	1.2480	1.6399	2.1012	1.6432	2.4265	.6314	.0824	.0532	.5541
HEM	41.4081	1.7434	1.2761	1.5883	1.9041	1.5727	2.1723	.6703	.1183	.0021	.4879
KERRVILLE	41.4251	1.7051	1.1474	1.6250	2.0343	1.6250	2.4456	.6260	.0846	.0122	.6207
LAMPSAS	41.5018	1.7933	1.2426	1.6594	2.1146	1.6491	2.6048	.6303	.0838	.0181	.5334
MISSION	41.5972	1.7030	1.0500	1.6704	2.1019	1.6764	2.5102	.6593	.0949	.0477	.7130
PEARTON	41.6950	1.7047	1.2294	1.6033	1.8522	1.6033	2.4780	.6347	.0719	.0211	.5950
POST	41.7206	1.5556	1.0892	1.5886	1.7396	1.5630	2.1993	.5980	.0555	.0048	.6784
PIESZIU	41.7262	1.5598	1.0741	1.5734	2.0597	1.7202	2.4050	.6718	.1018	.1018	.6622
RIVERSIDE	41.7300	1.7040	1.1250	1.6200	2.0947	1.6743	2.1214	.6013	.0861	.0514	.4999
STERLING CITY	41.8030	1.6947	1.2499	1.6910	2.0229	1.6910	2.2861	.6523	.0913	.1004	.5175
VALLEY JUNCTION	41.9280	1.7262	1.7049	1.7266	2.3385	1.7266	2.9561	.6821	.0913	.0373	.5373
VEGA	41.9330	1.7686	1.1407	1.6927	2.0362	1.6710	2.7290	.6109	.0511	.0367	.6203
WEATHERFORD	41.9532	1.6511	1.5664	1.7303	2.1963	1.7074	2.8467	.6257	.0773	.0222	.5889
DESERT	42.2101	1.6972	1.1303	1.5986	1.7494	1.5894	2.4212	.6535	.0856	.0306	.5717
FORT DUCHESSNE	42.3086	1.7011	1.1491	1.6201	2.1146	1.6201	2.4450	.6509	.0849	.0159	.6243
HIAWATHA	42.3966	1.7379	1.2523	1.7646	2.3437	1.7646	2.4242	.6290	.0840	.1155	.4879
KANAW POWER HOUSE	42.4508	1.6525	1.0164	1.5023	1.5833	1.5043	2.1920	.6507	.0765	.0103	.5769
LOA	42.5148	1.6993	1.1984	1.6399	2.0106	1.6399	2.6541	.5917	.0807	.0608	.6908
WILFORD #8 APT.	42.5654	1.6040	1.0799	1.5573	1.7333	1.5529	2.1972	.6526	.0853	.0451	.5157
RICHMOND	42.7271	1.8367	1.5372	1.6711	2.0211	1.6857	2.2818	.7090	.1107	.0668	.3946

STATION NAME	STATION NO.	EN ₂	Var N ₂	ES ₁	Var S ₁	ES ₂	Var S ₂	EI ₁	Var I ₁	EI ₂	Var I ₂
SPANISH FORK POWER HOU	42.8119	1.8276	1.5125	1.6421	1.9561	1.6370	2.2518	.7038	.1131	.8957	.4140
TUOELE	42.8771	1.7720	1.3980	1.5597	1.7078	1.5597	2.1247	.7066	.1121	.8802	.4454
BROOKLYN	45.0917	1.9176	1.7297	1.7060	2.1019	1.6963	2.3443	.7238	.1142	.8846	.3579
CEDAR LAKE	45.1223	1.6821	1.6052	1.5554	1.7601	1.5513	2.1842	.7214	.1127	.8331	.4144
CHELAN	45.1350	1.7606	1.3392	1.7040	2.1173	1.7040	2.5353	.6446	.0698	.9679	.5410
CULFAX 1 NW	45.1586	1.7188	1.2354	1.6105	2.0232	1.6105	2.3961	.6209	.0900	.9370	.5750
GULDENDALE	45.3222	1.6690	1.1166	1.6131	1.9297	1.6131	2.2614	.6310	.0815	.9665	.5589
HATTON 8 E	45.3546	1.7548	1.3246	1.6350	2.0009	1.6304	2.5253	.6279	.0812	.9291	.5967
LUNGVIEW	45.4769	1.8879	1.0761	1.6273	1.3421	1.6202	2.1717	.7467	.1240	.8582	.3553
NEWHALEM	45.5840	1.9238	1.7772	1.6594	1.9597	1.6518	2.2241	.7471	.1341	.8586	.3365
RIMROCK TETON DAM	45.7038	1.8116	1.4703	1.5768	2.0798	1.6768	2.6436	.6399	.0770	.9256	.5593
SEDRU WULLEY 1 E	45.7507	1.8095	1.4849	1.5906	1.8795	1.5906	2.1393	.7057	.1357	.8790	.4079
SHELTON	45.7584	1.7391	1.2856	1.5286	1.6484	1.5286	2.1904	.7031	.1177	.8789	.4208
SUNNYSIDE	45.8207	1.7095	1.2129	1.7171	2.1906	1.7131	2.4789	.6276	.0678	.10021	.5574
TATOOSH ISLAND WB	45.8332	1.8222	1.4783	1.6239	1.9599	1.6207	2.1934	.7063	.1410	.8894	.4038
WINTHROP 1 WSW	45.9376	1.8476	1.5561	1.7220	2.1331	1.7220	2.6578	.6599	.0695	.9320	.5086
BUFFALO											

Key Words: Run-Length, Run-Sum, Run-Intensity, Gamma and Normal Distributions, Moments

Abstract: Three quantitative measures are introduced for the concepts of "surplus" and "deficit" in hydrologic series. These are: run-length, run-sum and run-intensity. Positive and negative runs of a series are defined in terms of a fixed value, say c , of the variable under consideration, namely precipitation. The distribution function, moments, and other statistical properties of the three variables, run-length, run-sum, and run-intensity, are obtained analytically under the following alternative assumptions on the sequence of annual precipitations:

1. It is independent and normally distributed.
2. It is independent and gamma distributed.

For monthly precipitations, z_t , the series was first standardized by the transformation

$$x_t = \frac{z_t - \mu_t}{\sigma_t},$$

where t is of the form $t = 12(n-1) + \tau$, $\tau = 1, 2, \dots, 12$, $n = 1, 2, \dots$, and where μ_t and σ_t are mean and standard deviation of the series corresponding to the month τ . Calling " $x_t \leq c$ " as state "0" and " $x_t > c$ " as state "1", the series is then analyzed as a two-state Markov chain with stationary transition probabilities.

Annual precipitation from 27 stations in Colorado, and monthly precipitation from 219 stations in the Western United States are analyzed.

References: Jose Llamas and M. M. Siddiqui, Colorado State University Hydrology Paper No. 33 (May 1969), "Runs of Precipitation Series."

Key Words: Run-Length, Run-Sum, Run-Intensity, Gamma and Normal Distributions, Moments

Abstract: Three quantitative measures are introduced for the concepts of "surplus" and "deficit" in hydrologic series. These are: run-length, run-sum and run-intensity. Positive and negative runs of a series are defined in terms of a fixed value, say c , of the variable under consideration, namely precipitation. The distribution function, moments, and other statistical properties of the three variables, run-length, run-sum, and run-intensity, are obtained analytically under the following alternative assumptions on the sequence of annual precipitations:

1. It is independent and normally distributed.
2. It is independent and gamma distributed.

For monthly precipitations, z_t , the series was first standardized by the transformation

$$x_t = \frac{z_t - \mu_t}{\sigma_t},$$

where t is of the form $t = 12(n-1) + \tau$, $\tau = 1, 2, \dots, 12$, $n = 1, 2, \dots$, and where μ_t and σ_t are mean and standard deviation of the series corresponding to the month τ . Calling " $x_t \leq c$ " as state "0" and " $x_t > c$ " as state "1", the series is then analyzed as a two-state Markov chain with stationary transition probabilities.

Annual precipitation from 27 stations in Colorado, and monthly precipitation from 219 stations in the Western United States are analyzed.

References: Jose Llamas and M. M. Siddiqui, Colorado State University Hydrology Paper No. 33 (May 1969), "Runs of Precipitation Series."

Key Words: Run-Length, Run-Sum, Run-Intensity, Gamma and Normal Distributions, Moments

Abstract: Three quantitative measures are introduced for the concepts of "surplus" and "deficit" in hydrologic series. These are: run-length, run-sum and run-intensity. Positive and negative runs of a series are defined in terms of a fixed value, say c , of the variable under consideration, namely precipitation. The distribution function, moments, and other statistical properties of the three variables, run-length, run-sum, and run-intensity, are obtained analytically under the following alternative assumptions on the sequence of annual precipitations:

1. It is independent and normally distributed.
2. It is independent and gamma distributed.

For monthly precipitations, z_t , the series was first standardized by the transformation

$$x_t = \frac{z_t - \mu_t}{\sigma_t},$$

where t is of the form $t = 12(n-1) + \tau$, $\tau = 1, 2, \dots, 12$, $n = 1, 2, \dots$, and where μ_t and σ_t are mean and standard deviation of the series corresponding to the month τ . Calling " $x_t \leq c$ " as state "0" and " $x_t > c$ " as state "1", the series is then analyzed as a two-state Markov chain with stationary transition probabilities.

Annual precipitation from 27 stations in Colorado, and monthly precipitation from 219 stations in the Western United States are analyzed.

References: Jose Llamas and M. M. Siddiqui, Colorado State University Hydrology Paper No. 33 (May 1969), "Runs of Precipitation Series."

Key Words: Run-Length, Run-Sum, Run-Intensity, Gamma and Normal Distributions, Moments

Abstract: Three quantitative measures are introduced for the concepts of "surplus" and "deficit" in hydrologic series. These are: run-length, run-sum and run-intensity. Positive and negative runs of a series are defined in terms of a fixed value, say c , of the variable under consideration, namely precipitation. The distribution function, moments, and other statistical properties of the three variables, run-length, run-sum, and run-intensity, are obtained analytically under the following alternative assumptions on the sequence of annual precipitations:

1. It is independent and normally distributed.
2. It is independent and gamma distributed.

For monthly precipitations, z_t , the series was first standardized by the transformation

$$x_t = \frac{z_t - \mu_t}{\sigma_t},$$

where t is of the form $t = 12(n-1) + \tau$, $\tau = 1, 2, \dots, 12$, $n = 1, 2, \dots$, and where μ_t and σ_t are mean and standard deviation of the series corresponding to the month τ . Calling " $x_t \leq c$ " as state "0" and " $x_t > c$ " as state "1", the series is then analyzed as a two-state Markov chain with stationary transition probabilities.

Annual precipitation from 27 stations in Colorado, and monthly precipitation from 219 stations in the Western United States are analyzed.

References: Jose Llamas and M. M. Siddiqui, Colorado State University Hydrology Paper No. 33 (May 1969), "Runs of Precipitation Series."

PREVIOUSLY PUBLISHED PAPERS

Colorado State University Hydrology Papers

- No. 25 "An Experimental Rainfall-Runoff Facility," by W. T. Dickinson, M. E. Holland and G. L. Smith, September 1967.
- No. 26 "The Investigation of Relationship Between Hydrologic Time Series and Sun Spot Numbers," by Ignacio Rodriguez-Iturbe and Vujica Yevjevich, April 1968.
- No. 27 "Diffusion of Entrapped Gas From Porous Media," by Kenneth M. Adam and Arthur T. Corey, April 1968.
- No. 28 "Sampling Bacteria in a Mountain Stream," by Samuel H. Kunkle and James R. Meimann, March 1968.
- No. 29 "Estimating Design Floods from Extreme Rainfall," by Frederick C. Bell, July 1968.
- No. 30 "Conservation of Ground Water by Gravel Mulches," by A. T. Corey and W. D. Kemper, May 1968.
- No. 31 "Effects of Truncation on Dependence in Hydrologic Time Series," by Rezaul Karim Bhuiya and Vujica Yevjevich, November 1968.
- No. 32 "Properties of Non-Homogeneous Hydrologic Series," by V. Yevjevich and R. I. Jeng., April 1969.

Colorado State University Fluid Mechanics Papers

- No. 4 "Experiment on Wind Generated Waves on the Water Surface of a Laboratory Channel," by E. J. Plate and C. S. Yang, February 1966.
- No. 5 "Investigations of the Thermally Stratified Boundary Layer," by E. J. Plate and C. W. Lin, February 1966.
- No. 6 "Atmospheric Diffusion in the Earth's Boundary Layer--Diffusion in the Vertical Direction and Effects of the Thermal Stratification," by Shozo Ito, February 1966.

Colorado State University Hydraulics Papers

- No. 1 "Design of Conveyance Channels in Alluvial Materials," by D. B. Simons, March 1966.
- No. 2 "Diffusion of Slot Jets with Finite Orifice Length-Width Ratios," by V. Yevjevich, March 1966.
- No. 3 "Dispersion of Mass in Open-Channel Flow," by William W. Sayre, February 1968.